

# The effect of natural disturbances on forest biodiversity: an ecological synthesis

Mari-Liis Viljur<sup>1,\*</sup> , Scott R. Abella<sup>2</sup> , Martin Adámek<sup>3,4</sup> ,  
Janderson Batista Rodrigues Alencar<sup>5</sup> , Nicholas A. Barber<sup>6</sup> , Burkhard Beudert<sup>7</sup> ,  
Laura A. Burkle<sup>8</sup> , Luciano Cagnolo<sup>9</sup> , Brent R. Campos<sup>10</sup> , Anne Chao<sup>11</sup> ,  
Brahim Chergui<sup>12</sup> , Chang-Yong Choi<sup>13</sup> , Daniel F. R. Cleary<sup>14</sup> ,  
Thomas Seth Davis<sup>15</sup> , Yanus A. Dechnik-Vázquez<sup>16</sup> , William M. Downing<sup>17</sup> ,  
Andrés Fuentes-Ramírez<sup>18,19</sup> , Kamal J. K. Gandhi<sup>20</sup> , Catherine Gehring<sup>21</sup> ,  
Kostadin B. Georgiev<sup>1</sup> , Mark Gimbutas<sup>22</sup> , Konstantin B. Gongalsky<sup>23</sup> ,  
Anastasiya Y. Gorbunova<sup>23</sup> , Cathryn H. Greenberg<sup>24</sup> , Kristoffer Hylander<sup>25</sup> ,  
Erik S. Jules<sup>26</sup> , Daniil I. Korobushkin<sup>23</sup> , Kajar Köster<sup>27</sup> , Valerie Kurth<sup>28</sup> ,  
Joseph Drew Lanham<sup>29</sup> , Maria Lazarina<sup>30</sup> , Alexandro B. Leverkus<sup>31</sup> ,  
David Lindenmayer<sup>32</sup> , Daniel Magnabosco Marra<sup>33</sup> , Pablo Martín-Pinto<sup>34</sup> ,  
Jorge A. Meave<sup>35</sup> , Marco Moretti<sup>36</sup> , Hyun-Young Nam<sup>37</sup> , Martin K. Obrist<sup>36</sup> ,  
Theodora Petanidou<sup>30</sup> , Pere Pons<sup>38</sup> , Simon G. Potts<sup>39</sup> , Irina B. Rapoport<sup>40</sup> ,  
Paul R. Rhoades<sup>41</sup> , Clark Richter<sup>42</sup> , Ruslan A. Saifutdinov<sup>23</sup> , Nathan J. Sanders<sup>43</sup> ,  
Xavier Santos<sup>44</sup> , Zachary Steel<sup>45</sup> , Julia Tavella<sup>9,46</sup> , Clara Wendenburg<sup>38</sup> ,  
Beat Wermelinger<sup>47</sup> , Andrey S. Zaitsev<sup>23</sup>  and Simon Thorn<sup>1,48</sup> 

<sup>1</sup>Field Station Fabrikschleichach, Department of Animal Ecology and Tropical Biology (Zoology III), Julius Maximilians University Würzburg, Glashüttenstraße 5, 96181, Rauhenbrunn, Germany

<sup>2</sup>School of Life Sciences, University of Nevada Las Vegas, 4505 S. Maryland Parkway, Las Vegas, NV, 89154-4004, USA

<sup>3</sup>Department of GIS and Remote Sensing, Institute of Botany of the CAS, Příhonice, Czech Republic

<sup>4</sup>Department of Botany, Faculty of Science, Charles University, Benátská 2, CZ—128 01, Praha 2, Czech Republic

<sup>5</sup>Instituto Nacional de Pesquisas da Amazônia (INPA), Programa de pós-graduação em Ciências Biológicas (Entomologia), Manaus, AM, 0000-0001-9482-7866, Brazil

<sup>6</sup>Department of Biology, San Diego State University, 5500 Campanile Drive, San Diego, CA, 92182-4614, USA

<sup>7</sup>Bavarian Forest National Park, 94481, Grafenau, Germany

<sup>8</sup>Department of Ecology, Montana State University, Bozeman, MT, 59717, USA

<sup>9</sup>Instituto Multidisciplinario de Biología Vegetal (IMBIV), CONICET—Universidad Nacional de Córdoba, Vélez Sarsfield 1611, 5000, Córdoba, Argentina

<sup>10</sup>Point Blue Conservation Science, Petaluma, CA, 94954, USA

<sup>11</sup>Institute of Statistics, National Tsing Hua University, Hsin-Chu, 30043, Taiwan

<sup>12</sup>LESCB URL-CNRST N° 18, FS, Abdelmalek Essaadi University, Tetouan, Morocco

<sup>13</sup>Department of Agriculture, Forestry, and Bioresources, Seoul National University, Seoul, 08826, Republic of Korea

<sup>14</sup>CESAM and Department of Biology, University of Aveiro, 3810-193, Aveiro, Portugal

<sup>15</sup>Forest & Rangeland Stewardship, Warner College of Natural Resources, Colorado State University, Fort Collins, CO, 80523, USA

<sup>16</sup>Estudios Ambientales, Centro de Anteproyectos del Golfo, Comisión Federal de Electricidad, Nueva Era, Boca del Río, Veracruz, C.P. 94295, Mexico

<sup>17</sup>Department of Forest Ecosystems and Society, College of Forestry, Oregon State University, Corvallis, OR, 97331, USA

\* Author for correspondence (Tel.: +372 53 930 524; E-mail: [mariliisviljur@gmail.com](mailto:mariliisviljur@gmail.com)).

- <sup>18</sup>Laboratorio de Biometría, Departamento de Ciencias Forestales, Facultad de Ciencias Agropecuarias y Forestales, Universidad de La Frontera, Temuco, Chile
- <sup>19</sup>Centro Nacional de Excelencia para la Industria de la Madera (CENAMAD), Pontificia Universidad Católica de Chile, Santiago, Chile
- <sup>20</sup>D.B. Warnell School of Forestry and Natural Resources, University of Georgia, Athens, GA, 30602, USA
- <sup>21</sup>Department of Biological Sciences, Northern Arizona University, Flagstaff, AZ, 86011, USA
- <sup>22</sup>Institute of Mathematics and Statistics, University of Tartu, Narva mnt. 18, 51009, Tartu, Estonia
- <sup>23</sup>Severtsov Institute of Ecology and Evolution, Russian Academy of Sciences, Leninskii pr. 33, Moscow, 119071, Russia
- <sup>24</sup>USDA Forest Service, Southern Research Station, Bent Creek Experimental Forest, 1577 Brevard Road, Asheville, NC, 28806, USA
- <sup>25</sup>Department of Ecology, Environment and Plant Science, Stockholm University, SE-106 91, Stockholm, Sweden
- <sup>26</sup>Department of Biological Sciences, Humboldt State University, Arcata, CA, 95521, USA
- <sup>27</sup>Department of Environmental and Biological Sciences, Faculty of Sciences and Forestry, University of Eastern Finland, PL 111, 80101, Joensuu, Finland
- <sup>28</sup>Montana Department of Natural Resources and Conservation, Helena, MT, 59601, USA
- <sup>29</sup>Department of Forest Resources, Clemson University, 261 Lehotsky Hall, Clemson, SC, 29634, USA
- <sup>30</sup>Laboratory of Biogeography & Ecology, Department of Geography, University of the Aegean, University Hill, GR-81100, Mytilene, Greece
- <sup>31</sup>Department of Ecology, University of Granada, 18010, Granada, Spain
- <sup>32</sup>Fenner School of Environment and Society, The Australian National University, Canberra, ACT, Australia
- <sup>33</sup>Max-Planck-Institut für Biogeochemie, Hans-Knöll-Str 10, 07745, Jena, Germany
- <sup>34</sup>Sustainable Forest Management Research Institute, University of Valladolid, Avda, Madrid, Palencia, Spain
- <sup>35</sup>Departamento de Ecología y Recursos Naturales, Facultad de Ciencias, Universidad Nacional Autónoma de México, Coyoacán, Mexico City, 04510, Mexico
- <sup>36</sup>Swiss Federal Institute for Forest, Snow and Landscape Research WSL, Biodiversity and Conservation Biology, Zürcherstrasse 111, CH-8903, Birmensdorf, Switzerland
- <sup>37</sup>School of Biological Sciences, Seoul National University, Seoul, 08826, Republic of Korea
- <sup>38</sup>Departament de Ciències Ambientals, University of Girona, Campus Montilivi, 17003, Girona, Catalonia, Spain
- <sup>39</sup>Centre for Agri-Environmental Research, School of Agriculture, Policy and Development, University of Reading, Earley Gate, Reading, RG6 6AR, UK
- <sup>40</sup>Tembotov Institute of Ecology of Mountain Territories, Russian Academy of Sciences, I. Armand, 37a, Nalchik, Russia
- <sup>41</sup>Idaho State Department of Agriculture, Coeur d'Alene, ID 83854, USA
- <sup>42</sup>Science Department, Staten Island Academy, Staten Island, NY, USA
- <sup>43</sup>Department of Ecology and Evolutionary Biology, University of Michigan, 1105 North University Ave, Biological Sciences Building, Ann Arbor, MI, 48109-1085, USA
- <sup>44</sup>CIBIO-InBIO, Universidade do Porto, Campus de Vairão, Rua Padre Armando Quintas, 4485-661, Vairão, Portugal
- <sup>45</sup>Department of Environmental Science, Policy and Management, University of California, Berkeley, CA, 94720, USA
- <sup>46</sup>Facultad de Agronomía, Cátedra de Botánica General, Universidad de Buenos Aires, Buenos Aires, Argentina
- <sup>47</sup>Swiss Federal Institute for Forest, Snow and Landscape Research WSL, Forest Health and Biotic Interactions-Forest Entomology, Zürcherstrasse 111, CH-8903, Birmensdorf, Switzerland
- <sup>48</sup>Biology Centre of the Czech Academy of Sciences, Institute of Entomology, Branišovská 1160/31, 37005, České Budějovice, Czech Republic

## ABSTRACT

Disturbances alter biodiversity *via* their specific characteristics, including severity and extent in the landscape, which act at different temporal and spatial scales. Biodiversity response to disturbance also depends on the community characteristics and habitat requirements of species. Untangling the mechanistic interplay of these factors has guided disturbance ecology for decades, generating mixed scientific evidence of biodiversity responses to disturbance. Understanding the impact of natural disturbances on biodiversity is increasingly important due to human-induced changes in natural disturbance regimes. In many areas, major natural forest disturbances, such as wildfires, windstorms, and insect outbreaks, are becoming more frequent, intense, severe, and widespread due to climate change and land-use change. Conversely, the suppression of natural disturbances threatens disturbance-dependent biota. Using a meta-analytic approach, we analysed a global data set (with most sampling concentrated in temperate and boreal secondary forests) of species assemblages of 26 taxonomic groups, including plants, animals, and fungi collected from forests affected by wildfires, windstorms, and insect outbreaks. The overall effect of natural disturbances on  $\alpha$ -diversity did not differ significantly from zero, but some taxonomic groups responded positively to disturbance, while others tended to respond negatively.

Disturbance was beneficial for taxonomic groups preferring conditions associated with open canopies (e.g. hymenopterans and hoverflies), whereas ground-dwelling groups and/or groups typically associated with shady conditions (e.g. epigeic lichens and mycorrhizal fungi) were more likely to be negatively impacted by disturbance. Across all taxonomic groups, the highest  $\alpha$ -diversity in disturbed forest patches occurred under moderate disturbance severity, i.e. with approximately 55% of trees killed by disturbance. We further extended our meta-analysis by applying a unified diversity concept based on Hill numbers to estimate  $\alpha$ -diversity changes in different taxonomic groups across a gradient of disturbance severity measured at the stand scale and incorporating other disturbance features. We found that disturbance severity negatively affected diversity for Hill number  $q = 0$  but not for  $q = 1$  and  $q = 2$ , indicating that diversity–disturbance relationships are shaped by species relative abundances. Our synthesis of  $\alpha$ -diversity was extended by a synthesis of disturbance-induced change in species assemblages, and revealed that disturbance changes the  $\beta$ -diversity of multiple taxonomic groups, including some groups that were not affected at the  $\alpha$ -diversity level (birds and woody plants). Finally, we used mixed rarefaction/extrapolation to estimate biodiversity change as a function of the proportion of forests that were disturbed, i.e. the disturbance extent measured at the landscape scale. The comparison of intact and naturally disturbed forests revealed that both types of forests provide habitat for unique species assemblages, whereas species diversity in the mixture of disturbed and undisturbed forests peaked at intermediate values of disturbance extent in the simulated landscape. Hence, the relationship between  $\alpha$ -diversity and disturbance severity in disturbed forest stands was strikingly similar to the relationship between species richness and disturbance extent in a landscape consisting of both disturbed and undisturbed forest habitats. This result suggests that both moderate disturbance severity and moderate disturbance extent support the highest levels of biodiversity in contemporary forest landscapes.

**Key words:** natural disturbance, diversity–disturbance relationship, disturbance severity, disturbance extent, intermediate disturbance hypothesis, forest communities,  $\alpha$ -diversity,  $\beta$ -diversity.

## CONTENTS

|   |      |
|---|------|
| I. Introduction   | 1932 |
| II. Materials and methods   | 1934 |
| (1) Data collection   | 1934 |
| (2) Analysis of $\alpha$ -diversity                               | 1936 |
| (3) Analysis of $\beta$ -diversity                                | 1937 |
| (4) Analysis of mixture diversity                                 | 1937 |
| III. Results  | 1938 |
| (1) Effect of natural disturbance on $\alpha$ -diversity          | 1939 |
| (2) Effect of natural disturbance on $\beta$ -diversity           | 1940 |
| (3) Relationship between disturbance extent and mixture diversity | 1940 |
| IV. Discussion  | 1940 |
| (1) Overall effect and differences among taxonomic groups         | 1940 |
| (2) The effects of disturbance severity and extent                | 1941 |
| (3) The role of time since disturbance                            | 1942 |
| V. Conclusions  | 1943 |
| VI. Acknowledgements  | 1943 |
| VII. Data availability statement                                  | 1943 |
| VIII. References  | 1943 |
| IX. Supporting information  | 1946 |

## I. INTRODUCTION

Forest ecosystems host considerable terrestrial biodiversity, including many endangered species (Hill *et al.*, 2019). In turn, biodiversity plays a crucial role in forest ecosystem services, such as supporting the production of tree biomass, storing soil carbon, and providing goods and recreational services (Paquette & Messier, 2011; Gamfeldt *et al.*, 2013). Natural disturbances strongly influence biodiversity and can trigger major changes in forest communities (Swanson *et al.*, 2011; Lindenmayer *et al.*, 2019). The frequency, extent, intensity,

and severity of natural disturbances in forest landscapes – such as wildfires, windstorms, and insect outbreaks – is increasing in many parts of the world due to land-use modification and climate change (Seidl *et al.*, 2017; Sommerfeld *et al.*, 2018; Lindenmayer & Taylor, 2020; Collins *et al.*, 2021). Concurrently, the widespread suppression of natural disturbances can be detrimental to disturbance-dependent biota (Cumming, 2005; Hedwall & Mikusiński, 2016). Therefore, detailed knowledge about the response of forest communities to natural disturbances is essential.

Interest in the effects of disturbance and post-disturbance community succession is older than the scientific discipline of ecology itself (King, 1685; Clements, 1916). There are many theories that seek to explain how biodiversity responds to disturbance (e.g. Connell, 1978; Hastings, 1989; Wootton, 1998; reviewed by Pulsford, Lindenmayer & Driscoll, 2016). For instance, the popular but debated intermediate disturbance hypothesis (IDH) posits that diversity peaks at intermediate levels of disturbance frequency or intensity [Grime, 1973; Connell, 1978, but see Fox (2013) for a critique]. In addition to flaws in the theoretical underpinnings of the IDH, empirical studies have shown that relationships between species diversity and natural disturbance are often not hump-shaped but can have various shapes and strengths (Mackey & Currie, 2001; Bongers *et al.*, 2009). It has been hypothesised that disturbance intensity and frequency may have negligible effects on species diversity in communities where succession is driven by ecological processes other than competition (Chesson & Huntly, 1997). The diversity–disturbance relationship may differ depending on the organismal group, with hump-shaped responses for primary producers or sessile organisms, but non-hump-shaped responses for mobile organisms at higher trophic levels (Wootton, 1998). Disturbance effects on diversity are expected to be stronger following severe, infrequent disturbances as opposed to following moderate, more predictable disturbance events, such as regular flooding in riparian ecosystems (Reice, Wissmar & Naiman, 1990). Hence, while there is a plethora of theories and hypotheses, such as the IDH, there is no universal disturbance theory that could apply to different ecosystems, disturbance types, and taxonomic groups.

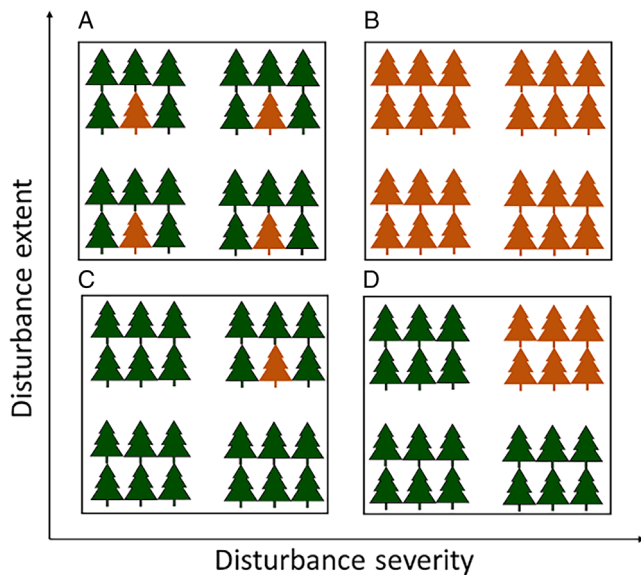
The characteristics of disturbance such as intensity, severity, and extent are critical in shaping diversity–disturbance relationships (Yeboah & Chen, 2016; Wernberg *et al.*, 2020). However, issues with defining terms as well as difficulties with quantifying disturbance characteristics in natural conditions at multiple spatial scales have impeded untangling the factors that drive the response of biodiversity to disturbance. Disturbance intensity, i.e. the strength of the disturbing force (e.g. the amount of heat produced by the wildfire; Keeley, 2009), can be difficult to quantify; thus this information is often missing in such studies (see Table 1 for definitions of terms). Disturbance severity can be applied more straightforwardly to compare the effects of different disturbance types. Disturbance severity can be quantified as the degree to which pre-disturbance vegetation has been damaged (Pickett & White, 1985). Specifically, to forests as the focus of our study, we considered the severity of a forest disturbance to be the proportion of trees in a stand with a destroyed canopy, which includes individuals of seeder species that are typically killed by disturbance and regenerate from seed as well as resprouter species that mostly survive and can resprout later (Pausas *et al.*, 2004). The term *disturbance severity* describes the impact of disturbance on the forest canopy at the scale of a disturbed forest stand. Conversely, we use the term *disturbance extent* to refer to the proportion of

Table 1. Glossary.

|                       |   |
|-----------------------|---|
| Disturbance intensity | Physical strength of the disturbing force, e.g. the amount of heat produced by a wildfire.  |
| Disturbance severity  | The effect of disturbance on an ecosystem, often quantified as the degree to which the pre-disturbance vegetation has been damaged. Herein we define disturbance severity as the proportion of trees with a completely killed canopy, including species typically killed by disturbance that regenerate from seed and resprouter species that mostly survive and can resprout after a disturbance event. The severity as defined here, serves as proxy for the strength of disturbance affecting demographic rates of above- and belowground organisms. |
| Disturbance extent    | The area affected by a disturbance. Herein we quantify disturbance extent as the proportion of disturbed area in a landscape consisting of disturbed and undisturbed forests.   |
| $\alpha$ -diversity   | Mean number of species sampled across a set of sampling localities.   |
| $\beta$ -diversity    | The extent of differentiation along habitat gradients according to Whittaker's (1960, 1972) definition. Several concepts and measures of $\beta$ -diversity exist in the literature. In the present meta-analysis, we use $\beta$ -diversity when we refer to differences in species composition between disturbed and undisturbed forests.   |
| Mixture diversity     | Diversity of a forest where a specified proportion of forest is disturbed, i.e. the diversity in a proportional mixture of disturbed and undisturbed forest.  |

disturbed habitat at the scale of a landscape where the landscape may have variable proportions of both disturbed and undisturbed forest stands (Fig. 1). Spatial scale of a landscape may vary from few hectares to several square kilometres depending on studied taxa and spatial extent of a particular study.

Disturbance characteristics can have varying effects on different components of biodiversity. Disturbances can increase, decrease or have no impact on the average species richness across a set of sampled localities, i.e.  $\alpha$ -diversity (Moretti, Obrist & Duelli, 2004; Ratchford *et al.*, 2005; Nelson, Halpern & Agee, 2008; Wernmeling *et al.*, 2017). However,  $\alpha$ -diversity does not reflect the disturbance-induced turnover in species composition, i.e.  $\beta$ -diversity. Extensive disturbances may create homogeneous disturbed areas and thus decrease  $\beta$ -diversity by homogenising species composition (Solar *et al.*, 2015). Conversely,  $\beta$ -diversity may increase if the variation in conditions is high following forest disturbance (Jones & Tingley, 2021). Variation in disturbance severity as well as in time since disturbance could create patchy landscapes with intermediate spatial extent of disturbed areas, resulting in higher biodiversity at the landscape level (Martin & Sapsis, 1992; Moretti *et al.*, 2002).



**Fig. 1.** Schematic representation of four hypothetical landscapes on the gradients of disturbance severity and extent. Each landscape consists of four forest stands. Dark green and brown trees represent undisturbed and disturbed parts of a stand, respectively. (A) Landscape where all stands are disturbed, but the disturbance severity within the stands is low; (B) landscape where all stands are severely disturbed; (C) landscape where most stands have not been affected by disturbance and the disturbance severity within a disturbed stand is low; (D) landscape where most stands have not been affected by disturbance, but the disturbance severity within a disturbed stand is high.

Here, we quantified the effects of natural disturbances on forest biodiversity on a global data set across 26 taxonomic groups, including plants, fungi, and animals from disturbed and undisturbed forests. We investigated the impacts of disturbance severity and extent separately and analysed their effects on different components of biodiversity. First, we predicted that natural disturbance changes  $\alpha$ -diversity but that the direction and magnitude of the effect depends on disturbance severity at the stand scale. To assess that expectation, we tested if natural disturbances influence  $\alpha$ -diversity of forest-dwelling species by comparing species richness in disturbed and undisturbed forests using a meta-analytic approach. We further tested if and how the response of  $\alpha$ -diversity depends on taxonomic group, disturbance severity, disturbance type, and time since disturbance. We predicted that several groups would benefit from disturbance, e.g. saproxylic organisms and groups associated with early successional stages, while groups preferring intact forest may be negatively affected. Additionally, we tested the effects of disturbance characteristics and taxonomic group on the disturbance-induced difference in  $\alpha$ -diversity by analysing species diversity calculated from species incidence matrices. Second, we hypothesised that natural disturbance induces changes in community composition and tested if the response of  $\beta$ -diversity, measured as the difference in species

composition between naturally disturbed and undisturbed forests, depends on taxonomic group and disturbance characteristics at the stand scale. Subsequently, we investigated the relationship between species diversity and disturbance extent at the landscape scale by simulating landscapes with variable proportions of disturbed and undisturbed patches. We tested the extent of disturbance at which forest biodiversity peaks, i.e. which level of disturbance extent corresponds to the maximal biodiversity considering both disturbed and undisturbed forest stands. For the latter, we used a mixed rarefaction/extrapolation approach to estimate the biodiversity change as a function of the proportion of disturbed forests on a landscape level (Chao *et al.*, 2019). This procedure allowed us to determine the disturbance extent that maximises diversity in a proportional mixture of disturbed and undisturbed habitats, i.e. maximal mixture diversity. We also tested if and how this level of disturbance extent at the landscape scale varies between taxonomic groups and depends on disturbance characteristics at the stand scale.

## II. MATERIALS AND METHODS

### (1) Data collection

We followed the guidelines for systematic literature reviews (Pullin & Stewart, 2006) to compile comparisons of species richness in naturally disturbed and undisturbed forests. We restricted our compilation to the three common types of natural forest disturbances: wildfires, windstorms, and insect outbreaks. To be included in our analyses, individual studies had to: (i) examine forests disturbed by wildfires, windstorms, or insect outbreaks; (ii) investigate forests affected by only one of the aforementioned disturbance types; (iii) compare disturbed and undisturbed control plots located in the same forest habitat type – compared disturbed and undisturbed plots had to be similar in size, surveyed with the same methods during the same study period and with the same sampling effort; (iv) provide spatially independent replicates of disturbed and undisturbed forests; (v) contain data about species assemblages (i.e. studies investigating populations of a single species were excluded); and (vi) be field-based case studies (i.e. simulation studies and reviews were excluded).

We screened the *Web of Science* electronic database using the following search string: (forest\* OR woodland\*) AND (disturb\* OR dieback OR wildfire OR fire OR windthrow OR storm OR pest OR ((insect\* OR beetle\*) AND (outbreak OR attack))) AND ('species richness' OR 'number\* of species' OR biodiversity). This resulted in >8000 articles (on 4 August 2020) from which we selected 338 articles after screening the titles and abstracts (see Fig. S1 for the decision workflow). In addition, we added four relevant studies, matching the same criteria, either found in the references of articles or articles that were suggested by authors. These 342 articles were further screened, and 51 studies matched the selection criteria based on the full texts.



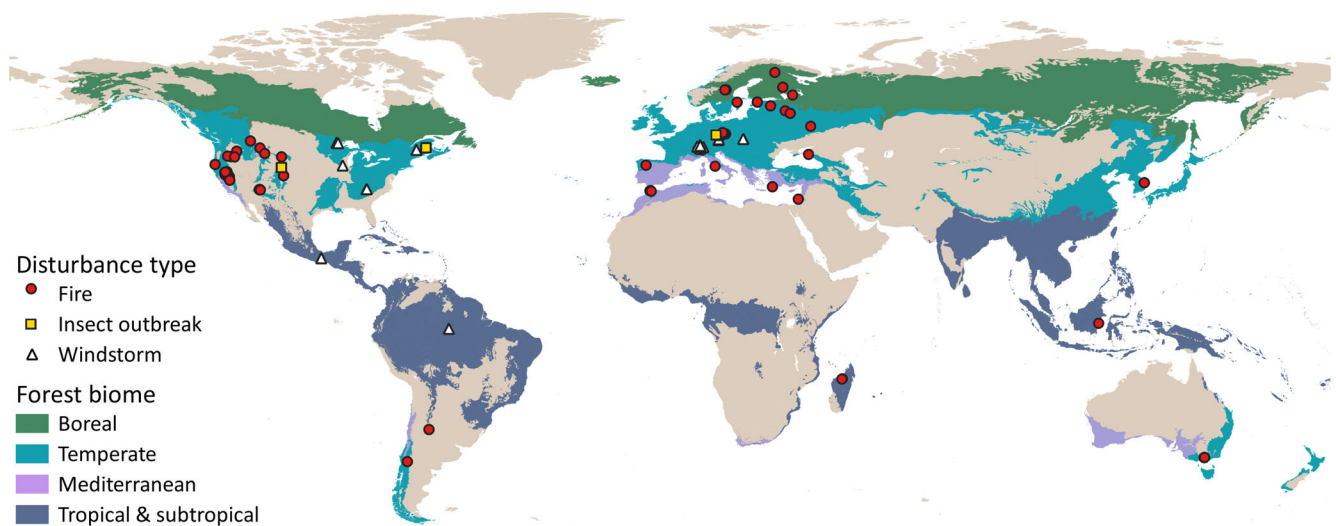
To ensure valid effect sizes, we assessed the quality of designs and excluded studies with pseudoreplication, i.e. studies where disturbed and undisturbed plots were not spatially interspersed (Hurlbert, 1984; Halme *et al.*, 2010). The spatial arrangement of plots in all studies was checked based on the description of methods and/or geographic coordinates. We contacted the authors of articles to provide data and to clarify the study designs when necessary.

Mean and standard deviation of the number of species in sampling units as well as the sample sizes in original studies were extracted from published texts, tables, and figures (we used <https://automeris.io/WebPlotDigitizer> for importing values from figures). In addition, we compiled information on covariates including disturbance type, i.e. wildfire, windstorm or insect outbreak, time since disturbance (in years), geographic location of the study area (latitude and longitude), disturbance severity, and taxonomic groups surveyed. Disturbance severity was quantified as the proportion of dead or fatally injured trees, per cent basal area loss or through qualitative indications (see Table S1 for information on all severity levels used in our study). In the few cases where only qualitative estimation was possible, we derived the following severity percentages as described in Appendix A in Leverkus *et al.* (2018b): low severity, 30%; low to moderate, 45%; moderate, 60%; moderate to high or mixed or variable, 75%; high, 90%; and severe, 100%. If the severity varied between study plots, we used the mean severity of plots for each study site, study year or severity category in the analyses (detailed information about methods for quantifying the severity is available in Table S1).

We gathered information describing the conditions of undisturbed control forests (summarised in Table S2). For that purpose, we asked authors to provide the following information: (i) if the control forests were primary or secondary forest; (ii) naturally grown or planted; (iii) the proportion

of non-native tree species; and (iv) the approximate age of the forest stand or a qualitative estimation of the successional phase. Alternatively, if authors could not be contacted, information was found from the source publications and/or the global map of forest classes (Schulze, Malek & Verburg, 2019). Most studies were conducted in secondary forests (68%). However, for 25% of the studies reported as primary forest, small-scale forest management activities such as selective logging and/or agricultural activities such as grazing/browsing within the last 100 years were evident or had possibly occurred. 84% of all control forests were naturally regrown and the proportions of non-native tree species were always reported as <20%, most often <2%; alternatively the proportion was qualitatively described as ‘none to very little’. Control forests in all study sites were typically mature, i.e. reached their typical local size making them ready to be harvested.

We compiled a database of species abundances/presences consisting of raw data underpinning the published studies. Authors of the respective studies provided raw data if not already available as supplementary material or in data repositories. Our final database consisted of 317 community matrices and included the following species groups: amphibians, ants, bats, birds, epigeic lichens, epigeic bryophytes, epigeic arachnids (spiders and harvestmen), hemipterans, herbaceous vascular plants, hoverflies (Syrphidae) and bee flies (Bombyliidae), hymenopterans (excluding ants), lacewings, lepidopterans (butterflies and moths), molluscs, mycorrhizal fungi, non-saproxyllic beetles, parasitoids (all recruited from Hymenoptera), reptiles, saproxyllic beetles, saproxyllic fungi, shrews, soil fauna (including soil-inhabiting springtails, oribatid mites and earthworms), tree-associated (epiphytic and epixylic) bryophytes, tree-associated (epixylic) lichens, and woody plants (trees and shrubs). Note that not all data from the meta-analysis of mean species richness were



**Fig. 2.** The locations of study sites ( $N = 70$ ) included in the meta-analysis. Classification of biomes is based on Olson *et al.* (2001).

underpinned by original raw data sets. One taxonomic group, stoneflies, was not supported by raw data and thus were included only in the meta-analysis of mean species richness. To account for differences among forests belonging to different biomes, each study site was classified as boreal, temperate, mediterranean, or subtropical/tropical using the classification of biomes in Olson *et al.* (2001) and the authors' specifications (see Fig. 2).

## (2) Analysis of $\alpha$ -diversity

We applied two approaches to analyse the effects of natural disturbances on  $\alpha$ -diversity: a meta-analysis of mean species richness and an analysis of species diversity quantified using Hill numbers. Hill numbers are a mathematically unified family of diversity measures that incorporate relative abundance and species richness. They are expressed in units of effective numbers of species, i.e. the number of equally abundant species that would be necessary to give the same value of a diversity measure (Chao *et al.*, 2014). For comparing mean numbers of species between naturally disturbed and undisturbed plots, we used the log response ratio (lnR) of species richness between the naturally disturbed plots (treatment group) and undisturbed plots (control group). The lnR describes the proportional difference in species richness between control and treatment groups. The natural logarithm transformation of the response ratio both linearises the metric, treating deviations in the denominator and the numerator as equal, and normalises its otherwise skewed distribution (Hedges, Gurevitch & Curtis, 1999). The meta-estimate of lnR across a set of studies can be back-transformed to per cent differences between treatment and control groups, providing an intuitively accessible effect size value.

We fitted two types of models. First, we conducted a random-effects meta-analysis to assess if the overall effect of natural disturbance on species richness differed from zero. For that purpose, we fitted only the intercept (i.e. the overall mean effect size measured as lnR) and random effects. To quantify heterogeneity in this model, we used the  $Q$  statistic (Viechtbauer, 2010). The identity of a case study and its geographic coordinates were included as random effects in the model to control for unmeasured specificities of sampling design and study area as well as for spatial and temporal autocorrelation. This means that multiple data points per study were included if studies examined multiple taxonomic groups, if surveys lasted for more than 1 year, or if they investigated multiple disturbance types (affecting different study plots) or different ranges of disturbance severities. To account for dependent effect sizes caused by the same control group being shared between different disturbed plots (differing either in disturbance severity or in time since disturbance), we calculated the corrected variance–covariance matrix and used this as a component of sampling variance meta-regression models (Lajeunesse, 2011).

Second, we examined if and how the effect of disturbance depends on taxonomic group, characteristics of disturbance, and the study system. For that purpose, we conducted a mixed-effects meta-analysis by adding the fixed moderators (i.e. effect modifiers) to the model described above. These fixed moderators included the taxonomic group, disturbance type, and forest biome as categorical variables. Disturbance severity and time since disturbance (years) were included as quantitative variables. Time since disturbance was log-transformed to improve normality. Both numeric variables were centred by subtracting the mean. We also included the quadratic terms of time since disturbance and disturbance severity to test the non-linear relationship with the lnR. We subtracted the intercept from the effect sizes (by including  $-1$  in the model formula). We obtained the significance of effect modifiers from  $Q$  tests covering all the parameters related to a given factor (Viechtbauer, 2010). In the final model, we used an omnibus test of moderators ( $Q_M$ ) to assess the null hypothesis that all coefficients except the intercept were simultaneously zero, and a further test ( $Q_E$ ) to assess the significance of residual heterogeneity. We considered all effects significant at  $P \leq 0.05$ .

We estimated species diversity for Hill numbers  $q = 0, 1$ , and  $2$  based on 95% sample coverage for disturbed and undisturbed plots, based on species incidence (Chao *et al.*, 2014). This procedure allowed us to compare observed levels of  $\alpha$ -diversity to levels of  $\alpha$ -diversity standardised by sampling effort. Furthermore, the statistical framework based on Hill numbers reveals the importance of species relative abundance in responding to disturbance, i.e. whether disturbance effects on  $\alpha$ -diversity were predominantly driven by rare, common or dominant species [Chao *et al.*, 2020; see Thorn *et al.*, 2020a and Georgiev *et al.*, 2020 for a similar application of the Hill framework]. We calculated the difference between estimated Hill numbers as lnR between naturally disturbed and undisturbed plots. We further used this as a response variable in the linear mixed-effects model including taxonomic group, disturbance type, and forest biome as categorical variables. Disturbance severity and time since disturbance (years) were included as quantitative variables. Time since disturbance was log-transformed to improve normality. Both numeric variables were centred by subtracting the mean. We also included the quadratic terms of time since disturbance and disturbance severity to test the non-linear relationship with the lnR. We subtracted the intercept from the effect sizes (by including  $-1$  in the model formula). The identity of a case study and study site were included as random effects.

All analyses were conducted in R 4.0.4 (R Core Team, 2021). We used the *rma.mv* function in the *metafor* package (Viechtbauer, 2010) for building and testing the multi-level linear mixed models. lnR values were weighted by the corresponding sampling variance within the statistical model. Hill numbers for specified sample coverage were calculated using the *estimateD* function in the *iNEXT* package (Hsieh, Ma & Chao, 2016).

### (3) Analysis of $\beta$ -diversity

We used original raw data matrices to examine changes in community composition induced by natural disturbance. A raw data matrix consisted of presences and absences of species in all study plots sampled in one case study in the same study site and during the same year since disturbance, for both disturbed and undisturbed forests. For testing differences in community composition between disturbed and undisturbed forests, we conducted permutational multivariate analysis of variance (PERMANOVA; Anderson, 2017). The PERMANOVA was based on matrices of Jaccard distances and performed using the function *adonis2* in the *vegan* package (Oksanen *et al.*, 2019). This analysis provides a pseudo *F*-value, based on 999 permutations, that quantifies the deviance from the null hypothesis (which states that composition is the same between disturbed and undisturbed forest), while simultaneously accounting for imbalanced study designs (McCordle & Anderson, 2001). Consequently, larger pseudo *F*-values correspond to larger changes in community composition resulting from natural disturbance. This pseudo *F*-value represents the standardised difference between communities in disturbed and undisturbed plots within a single presence–absence matrix. We restricted this analysis to those matrices that yielded pseudo *F*-values over the course of permutations. For this purpose, those matrices that generated fewer than 99 permutations were excluded. These restrictions resulted in a total of 209 matrices, which provided pseudo *F*-values for the analysis described below.

To test if natural disturbance affected community composition, we modelled pseudo *F*-values using linear mixed models with the *lmer* function in the *lme4* package (Bates *et al.*, 2015). To test if the association depended on the characteristics of the disturbance and the study system, we included the taxonomic group, disturbance type and forest biome as categorical predictors, and disturbance severity and time since disturbance as numerical covariates. For the analysis of  $\beta$ -diversity, we included the study identity and study site as random effects to control for possible differences among study sites and repeated measurements within one study site. We omitted the intercept from the model formula to determine if pseudo *F*-values differed significantly from zero. Therefore, significant changes in community composition due to natural disturbance were indicated by pseudo *F*-values significantly larger than zero.

### (4) Analysis of mixture diversity

We estimated species diversity associated with landscapes within which a given proportion of forest was affected by disturbance using a mixed rarefaction/extrapolation approach based on Hill numbers (Hill, 1973; Chao *et al.*, 2019; for an application see Thorn *et al.*, 2020b). The method extends classical rarefaction and extrapolation towards a proportional mixture of two rarefaction/extrapolation curves derived from two distinct species

assemblages (Chao *et al.*, 2014, 2019). The mixed rarefaction/extrapolation was based on plots surveyed in undisturbed forests (t1) and plots surveyed in corresponding disturbed forests (t2). When a proportion of undisturbed plots (e.g. t1a) are disturbed by a natural disturbance, it is equivalent to replacing these t1a plots with the same number of plots randomly selected from disturbed forests. Using a mixture of rarefaction and extrapolation, we analytically retrieved the species richness of mixed assemblages and assessed species compositions (i.e. the number of unique and shared species). In addition to this, mixed rarefaction/extrapolation allowed estimating the proportions of disturbed and undisturbed plots corresponding to certain values of species richness. These values of species richness include the numbers of species unique to disturbed or undisturbed forests as well as the numbers of species that are shared between disturbed and undisturbed forests. By summing the richness of unique species and shared species, the total landscape-level species richness of the proportional mixture can be calculated for each proportional mixture of disturbed and undisturbed forests.

The proportion of plots affected by disturbance can subsequently be used as a proxy for the proportion of forest area affected by disturbance (Chao *et al.*, 2019; Thorn *et al.*, 2020b). Mixed rarefaction/extrapolation is based on comparisons of plots randomly selected from any location of a study design and is independent of plot size and the number of plots within a respective study, as long as all plots within a study are of similar sizes. The independence from the sizes of the areas is particularly important as detailed information about the size of a disturbed area for each study year was not available. We plotted the mixture diversity as well as the numbers of species unique to disturbed and undisturbed forests against the proportion of disturbed forest and fitted curves using LOESS (local polynomial regression) by means of the basic *loess* function in R with the *family* argument set as ‘symmetric’ to obtain more robust curves, i.e. to reduce the impact of possible outliers. We further estimated the proportion of disturbed forest area that is associated with the maximal levels of mixture diversity for each individual species matrix (i.e. per each taxonomic group in a study site and particular year after disturbance). The analyses were conducted using the R code miNEXT (mixed iNterpolation/EXTrapolation, available at <https://github.com/AnneChao>).

We tested the effect of disturbance type, taxonomic group, forest biome, disturbance severity, and time since disturbance on the proportion of disturbed forest associated with the highest mixture diversity per data set. For that purpose, we used beta regression by means of the function *gam* with the *family* argument set to *betar* and logit-link function in the R-package *mgcv* (Wood, Pya & Saefken, 2016). We subtracted the intercept from the effect sizes (by including  $-1$  in the model formula) to evaluate if effects differed significantly from 50% of plots affected by disturbance. To make the coefficients of the model interpretable, we included

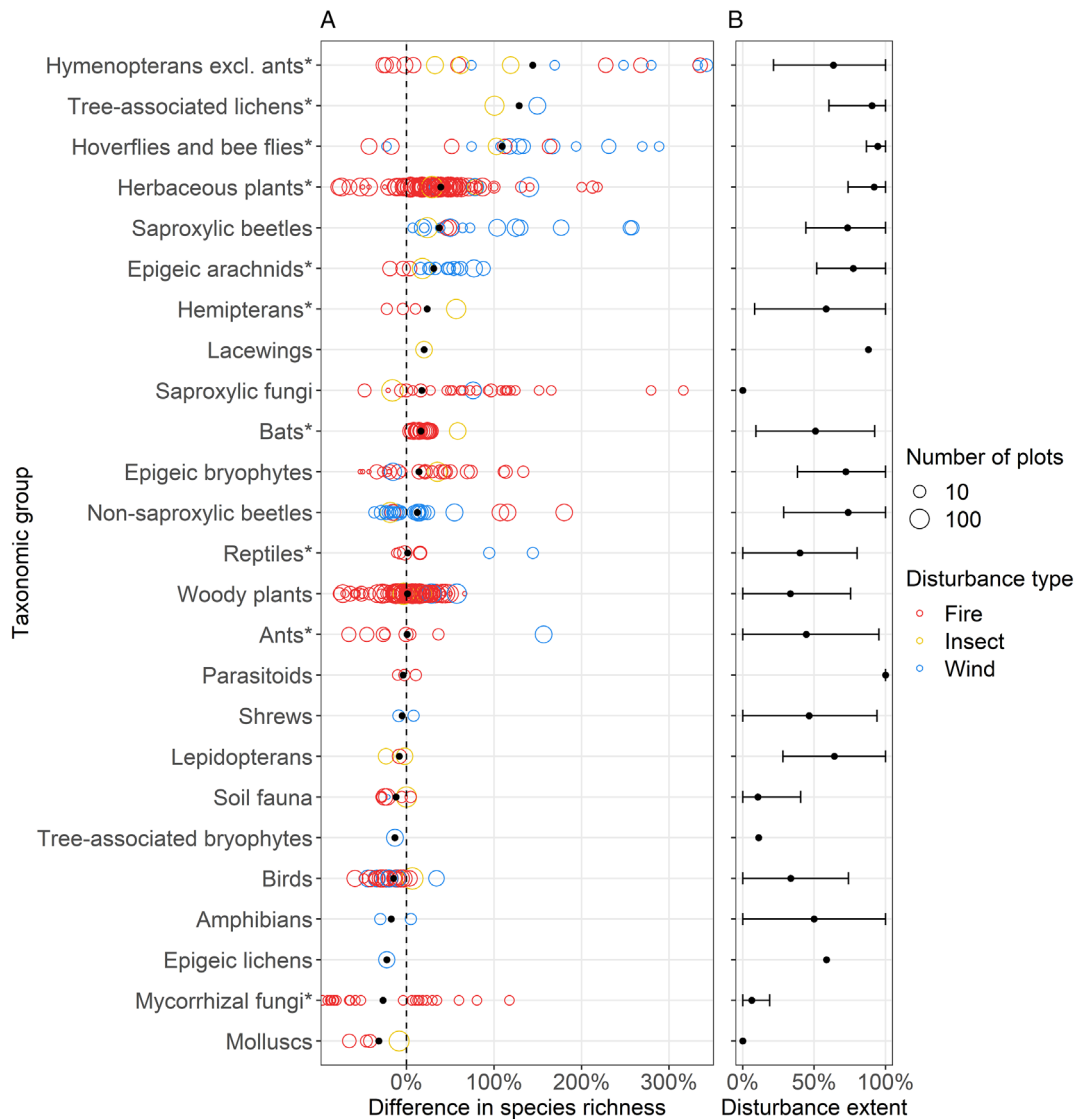


non-numerical covariates as random factors and centred the numerical covariates: disturbance severity and time since disturbance. We also controlled for study identity and study site by including these variables as random effects.

III. RESULTS

Our analysis of  $\alpha$ -diversity was based on 508 data points, with comparisons between disturbed and undisturbed forests in 70 different study sites derived from 51 studies (Fig. 2). Most studies were conducted in boreal and temperate biomes

(82%), with relatively few (18%) in mediterranean and tropical regions (Figs 2 and S2). Wildfires were the most commonly investigated disturbance (33 studies, 381 data points), followed by windstorms (15 studies, 98 data points) and insect outbreaks (4 studies, 29 data points). Among the 26 taxonomic groups, herbaceous and woody vascular plants were the most widely studied groups (13 studies, 98 data points and 11 studies, 97 data points, respectively), followed by non-saproxyllic beetles (9 studies, 28 data points). Early years after disturbance were studied most often – of the 508 data points, 377 covered 10 years or less, 94 represented 11–20 years and 37 data points addressed more than 20 years after disturbance. The numbers of individual studies



**Fig. 3.** Differences in  $\alpha$ -diversity between naturally disturbed and undisturbed forests categorised by taxonomic group. (A) Black dots represent weighted mean effect sizes transformed to percentage differences in species richness, coloured circles correspond to each data point. Asterisks against group names indicate significant effects ( $p < 0.05$ ) in the mixed-effects meta-analysis. (B) Proportions of disturbed forest that correspond to the maximal mixture diversity of each taxonomic group. Dots represent the means of data points and whiskers show respective standard deviations. Whiskers are missing when only one incidence matrix for a taxonomic group was available.

per range of years were 46, 16, and 10, respectively. In the analyses of  $\beta$ -diversity and mixture diversity, we used 209 and 297 full species-by-plot abundance matrices, respectively. These subsets were similar in the aforementioned aspects – covering primarily boreal and temperate biomes, wildfires, and vascular plants collected within 20 years since disturbance.

### (1) Effect of natural disturbance on $\alpha$ -diversity

Across all taxa, disturbance types, disturbance severities, forest biomes, and time after disturbance, the effect of natural disturbance on  $\alpha$ -diversity did not differ significantly from zero ( $z = -0.127$ ;  $p = 0.899$ ). There was high heterogeneity

Table 2. Results of mixed-effects meta-analysis of the effect of natural disturbances on  $\alpha$ -diversity. The response variable is the difference in mean species numbers between disturbed and undisturbed forests, quantified as  $\ln R$ .  $Q_M$  = omnibus test for all the parameters in the model;  $Q_E$  = test for residual heterogeneity.

| Tested moderator                    | $Q$     | df  | $p$              |
|-------------------------------------|---------|-----|------------------|
| Disturbance severity                | 195.30  | 1   | <b>&lt;0.001</b> |
| Disturbance severity <sup>2</sup>   | 98.31   | 1   | <b>&lt;0.001</b> |
| Time since disturbance              | 1.29    | 1   | 0.256            |
| Time since disturbance <sup>2</sup> | 41.51   | 1   | <b>&lt;0.001</b> |
| Taxonomic group                     | 862.71  | 26  | <b>&lt;0.001</b> |
| Disturbance type                    | 2.54    | 2   | 0.28             |
| Forest biome                        | 12.05   | 3   | <b>0.007</b>     |
| <b>Full model</b>                   |         |     |                  |
| $Q_M$                               | 1099.54 | 35  | <b>&lt;0.001</b> |
| $Q_E$                               | 3953.33 | 461 | <b>&lt;0.001</b> |

Bold values indicate significant effects ( $p < 0.05$ ).

among studies in this model ( $Q = 5915.39$ ; degrees of freedom [df] = 495;  $p < 0.001$ ), meaning that the effect of disturbance varied considerably among studies.

The mixed-effects meta-analysis showed that the impacts of natural disturbance on  $\alpha$ -diversity varied among taxonomic groups (Fig. 3; Table S3) and depended on disturbance severity and time since disturbance (Table 2). The relationships between  $\alpha$ -diversity and the quadratic terms of disturbance severity and time since disturbance were negative, indicating hump-shaped relationships with the standardised difference in mean species richness. The positive effect of disturbance on  $\alpha$ -diversity peaked at  $\sim 55\%$  disturbance severity and decreased thereafter (Fig. 4A). Time since disturbance moderated the effect of disturbance on  $\alpha$ -diversity by driving it towards a positive direction from about 2–3 years after the disturbance event and reaching a maximum around 10 years post-disturbance (Fig. 4B). Forest biome moderated disturbance effects: the impact was more

Table 3. Type III analysis of variance table of the results of the linear mixed-effects model with pseudo  $F$ -value as response variable. Pseudo  $F$  represents a disturbance-induced difference in species composition. Note that  $F$ -values in the table are  $F$ -statistics from the Type III analysis of variance (not pseudo  $F$ -values).

| Explanatory variable   | $F$  | df | $p$              |
|------------------------|------|----|------------------|
| Taxonomic group        | 2.97 | 23 | <b>&lt;0.001</b> |
| Disturbance type       | 0.84 | 2  | 0.445            |
| Forest biome           | 0.58 | 3  | 0.634            |
| Disturbance severity   | 2.51 | 1  | 0.116            |
| Time since disturbance | 1.11 | 1  | 0.295            |

Bold values indicate significant effects ( $p < 0.05$ ).

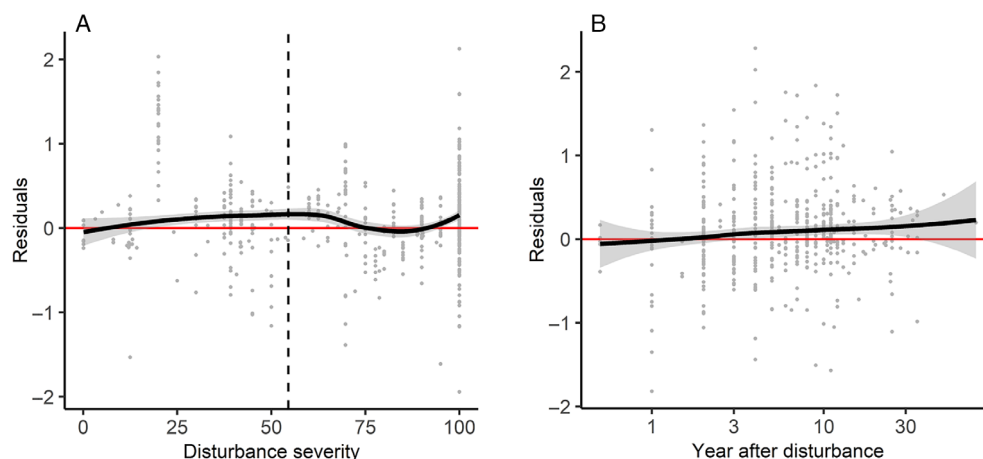
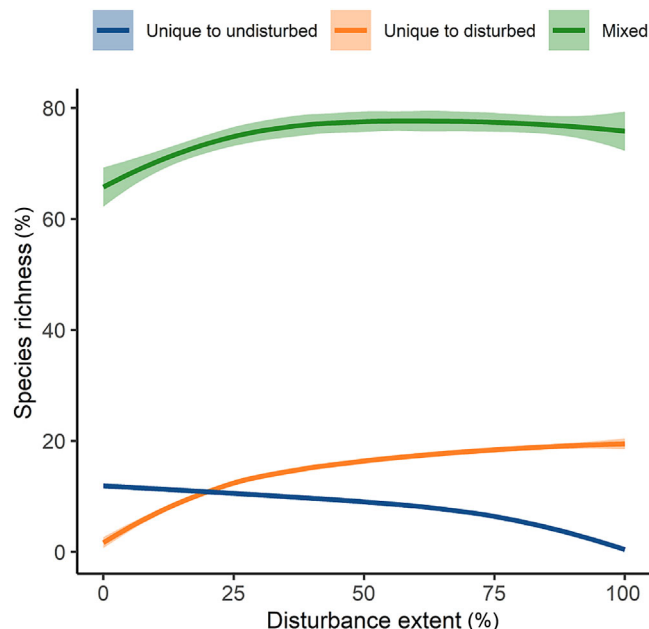


Fig. 4. The residuals of the models explaining the effect of natural disturbance on  $\alpha$ -diversity using mixed-effects meta-analysis plotted against (A) disturbance severity, and (B) time since disturbance (year is logarithmic and then back-transformed for easier readability). Descriptions of full models are provided in Section II.(2). Residuals in A result from the full model excluding disturbance severity. Residuals in B result from the full model excluding the time since disturbance. The curves and confidence intervals (shading) are generated by fitting the LOESS function. The dashed vertical line marks the maximal value of  $\alpha$ -diversity in disturbed plots that remained unexplained by other moderators except disturbance severity, i.e. the predicted value of disturbance severity that corresponds to the maximal  $\alpha$ -diversity.



**Fig. 5.** Response of species richness to the proportion of disturbed area in the forest landscape. The curves and confidence intervals (shading) are generated by fitting the LOESS function. Components of species richness: species unique to disturbed forests; species unique to undisturbed forests; a mixture of both groups of unique species and species shared between two habitat types, i.e. mixture diversity.

Table 4. Results of the beta-regression with the proportion of disturbed forest corresponding to highest mixture diversity as a response variable. Wald tests on terms were applied to test for significance.

| Explanatory variable   | $\chi^2$ | df | $p$          |
|------------------------|----------|----|--------------|
| Taxonomic group        | 49.908   | 25 | <b>0.002</b> |
| Disturbance type       | 2.743    | 2  | 0.077        |
| Forest biome           | 4.459    | 3  | <b>0.037</b> |
| Disturbance severity   | 1.372    | 1  | 0.242        |
| Time since disturbance | 5.285    | 1  | <b>0.022</b> |

Bold values indicate significant effects ( $p < 0.05$ ).

negative in mediterranean forests compared to temperate and boreal forests (Tables 2 and S3; Fig. S2). Disturbance type did not significantly moderate the effect of disturbance (Table 2).

Our analysis of disturbance effects on biodiversity quantified using Hill numbers ( $q = 0, 1, 2$ ) for 95% of sample coverage showed that the findings described above were largely driven by species relative abundance. We found that  $\ln R$  for  $q = 0$  was negatively associated with disturbance severity ( $z = -2.55$ ;  $p = 0.01$ ; Table S4).  $\ln R$  of all three Hill numbers depended significantly on the forest biome (Tables S4–S9; Figs S3–S5). There were significant positive effects of disturbance on  $\ln R$  for the following taxonomic groups: hymenopterans (excluding ants), hoverflies and bee flies, herbaceous plants, non-saproxyllic beetles (for  $q = 0$  only), and saproxyllic beetles.

(2) Effect of natural disturbance on  $\beta$ -diversity

Natural disturbance was associated with significant changes in the species compositions of six taxonomic groups (Table S10): epigeic arachnids ( $z = 2.07$ ;  $p = 0.038$ ), birds ( $z = 3.66$ ;  $p < 0.001$ ), tree-associated (epixylic) lichens ( $z = 2.93$ ;  $p = 0.003$ ), woody plants ( $z = 2.86$ ;  $p = 0.004$ ), herbaceous plants ( $z = 2.25$ ;  $p = 0.024$ ), and ants ( $z = 2.11$ ;  $p = 0.035$ ). Disturbance type, forest biome, disturbance severity, and time since disturbance did not significantly modulate the effects of disturbance on species composition (Table 3).

(3) Relationship between disturbance extent and mixture diversity

Mixture diversity (encompassing species unique to disturbed and undisturbed forest as well as shared species) peaked at ~60% of disturbed forest (Fig. 5) but varied considerably among taxa (Fig. 3B). Beta-regression showed that disturbance extent (i.e. the proportion of forest that was disturbed) corresponding to the highest mixture diversity differed significantly among taxonomic groups (Tables 4 and S11). The disturbance extent corresponding to the highest mixture diversity was significantly larger than 50% for hymenopterans (excl. ants) (at 77% of disturbed forest;  $p < 0.001$ ), hoverflies and bee flies (at 76% of disturbed forest;  $p = 0.004$ ) and reptiles (at 76% of disturbed forest;  $p = 0.034$ ). The disturbance extent corresponding to the highest mixture diversity was greater for longer times since disturbance (Tables 4 and S11).

IV. DISCUSSION

(1) Overall effect and differences among taxonomic groups

Our global meta-analysis (with most sampling concentrated in boreal and temperate biomes), showed that the overall effect of natural disturbance on forest  $\alpha$ -diversity did not differ significantly from zero. However, the responses varied significantly among taxonomic groups, and while many groups benefited from disturbance, several others were negatively impacted, possibly due to different disturbance response strategies (Moretti & Legg, 2009; Pausas, 2019). For example, flower-visiting hymenopterans and flies were among the taxa that benefited from natural disturbance, possibly because of the increased availability of nectar and pollen resources and warmer microclimatic conditions caused by enhanced canopy openness at early successional stages (Proctor *et al.*, 2012; Rodríguez & Kouki, 2017; Ramos-Fabiel *et al.*, 2018). Our finding is consistent with a meta-analysis on the effect of fire on pollinators that revealed positive responses of this group to wildfire (Carbone *et al.*, 2019). Conversely, ground-dwelling taxa and/or groups typically associated with shady conditions, such as epigeic lichens,

and mycorrhizal fungi, tended to be more often negatively affected by disturbances (Moretti *et al.*, 2004). Mycorrhizal fungi possibly also additionally suffered from the death of host trees poorly adapted to severe wildfires (Franco-Manchón *et al.*, 2019).

Previous studies have often reported positive multi-taxa responses to natural disturbances. In a previous meta-analysis investigating disturbance effects on temperate and boreal forest ecosystems, Thom & Seidl (2015) found that stand-replacing natural disturbances increased species richness by 36% on average. They concluded that the impact of disturbance was variable, with some studies in their data set reporting negative effects of disturbances on some elements of biodiversity, but the overall response of biodiversity to disturbance was positive (Thom & Seidl, 2015). The high variability in multi-taxa responses, including taxa identified by sequencing, resulting in the absence of an overall effect of disturbance, has been previously found in forests affected by bark beetles (Kortmann *et al.*, 2021). Another study on multi-taxa responses to bark beetle disturbance reported a positive overall effect on biodiversity (Beudert *et al.*, 2015). However, differences in the response to disturbance have not been tested previously for such a wide variety of taxonomic groups and ecosystems as in our meta-analysis. Our study shows that the heterogeneity of responses to disturbances in a large global data set prevails over possible directional effects in smaller subsets of taxa.

In the analysis of  $\alpha$ -diversity using Hill numbers and accounting for sample coverage, differences between diversities in disturbed and undisturbed forests for most taxonomic groups were non-significant. Natural disturbances significantly increased the diversity of hymenopterans, hoverflies (Syrphidae) and bee flies (Bombyliidae), saproxylic beetles, non-saproxylic beetles, and herbaceous plant species (for  $q = 0$ ). The list of positively affected taxonomic groups was almost identical for  $q = 1$  and  $q = 2$ , excluding only non-saproxylic beetles. The higher diversity of saproxylic beetles in disturbed compared to undisturbed plots is not surprising, as saproxylic beetles benefit from abundant dead wood and warm microclimates (Seibold *et al.*, 2016, 2021). It is possible that the higher diversity of herbaceous plants in disturbed plots due to open canopy conditions at least partly explains the high numbers of flower-visiting Hymenoptera and Diptera, as well as non-saproxylic beetle species. Bees, the most widely studied group within Hymenoptera, benefit from diverse plant communities, which provide a variety of nectar resources and support bees during longer vegetative periods (Rubene, Schroeder & Ranius, 2015; Rhoades *et al.*, 2018). A considerable number of non-saproxylic beetles are herbivores that may also benefit from higher plant diversity. The other large group of non-saproxylic beetles was composed of mainly carnivorous ground beetles (Carabidae) that may benefit from the abundance of other invertebrates as well as the variety of microhabitats created by natural disturbances (Barber & Widick, 2017). The lack of statistically significant disturbance effects for several other taxa in this additional analysis indicates the substantial role of species relative

abundances in shaping diversity–disturbance relationships of different taxonomic groups.

The effect of a natural disturbance on biodiversity may depend on the pre-disturbance condition of the forest (Donato, Campbell & Franklin, 2012; Lindenmayer *et al.*, 2019). In our synthesis, an undisturbed control forest provided a proxy for the pre-disturbance state of a respective disturbed forest. Forest management intensity is one of the important aspects that could potentially moderate the contrast in species richness between naturally disturbed and control forests. However, the vast majority of forests in our data set were relatively homogenous in management intensity and stand age, with most having experienced weak to moderate management activities and were in intermediate to late successional stages (summarised in Table S2). This prevented us from investigating the effect of pre-disturbance conditions on the outcomes of forest disturbance on biodiversity.

## (2) The effects of disturbance severity and extent

Our results suggest that moderate disturbance severity and moderate disturbance extent were associated with maximum overall levels of biodiversity at the stand and landscape scale, respectively. Our meta-analysis showed that the highest levels of  $\alpha$ -diversity in disturbed stands were associated with forests in which slightly more than half of the trees (~55%) had canopies destroyed by disturbance;  $\alpha$ -diversity declined when the severity increased further (see Table 1 for the definition of disturbance severity). The hump-shaped relationship found in our study is in accordance with the relationship predicted by the intermediate disturbance hypothesis (Connell, 1978). The hump-shaped relationship can be explained by potentially higher structural complexity in moderately disturbed stands contributing to higher diversity of microhabitat conditions and resources compared to more homogeneously disturbed or nearly undisturbed stands (Stein, Gerstner & Kreft, 2014; Senf *et al.*, 2020). That is, disturbances with intermediate severity may be associated with higher 'disturbance diversity' that can be related to higher biodiversity as proposed by the pyrodiversity–biodiversity hypothesis (Martin & Sapsis, 1992; Jones & Tingley, 2021; Steel *et al.*, 2021). Some case studies showing that intermediate fire severity can lead to higher pyrodiversity and consequently enhance biodiversity provide support for that explanation (Ponisio *et al.*, 2016; Lazarina *et al.*, 2019). However, the relationship between  $\alpha$ -diversity and disturbance severity was not hump-shaped when standardising for sampling effort. The effect was linear negative for  $q = 0$ , indicating that rare species are more likely to be negatively affected by severe disturbances and/or severely disturbed forests are dominated by few abundant disturbance-favoured species. This finding is in line with the results of case studies, demonstrating that rare species were most affected by post-disturbance management, regardless of taxonomic group (Thorn *et al.*, 2020a). This could be explained by the higher probability that rare species become extinct from a local site, while common and dominant species often have a higher probability of colonising a given site (Chao



*et al.*, 2019). While rare species are more likely to be affected by stochastic processes such as chance of colonisation, random extinction and ecological drift, common and dominant species are more likely to be affected by deterministic processes including environmental filtering, and biotic interactions (Chase & Myers, 2011).

The availability of habitat patches in different successional stages may also result in high biodiversity levels at the landscape scale (Fahrig, 2020). This explanation is also supported by our mixed rarefaction/extrapolation approach, which showed that the overall maximal mixture diversity corresponded to disturbance extents that were close to intermediate, indicating that disturbed and undisturbed patches were both present at similar amounts. Our mixed rarefaction/extrapolation approach showed that the mixture diversity peaked at 60% disturbed forest within the landscape. Similar to the relationship between  $\alpha$ -diversity and disturbance severity, the maximal mixture diversity corresponded to a disturbance extent that was moderate, but somewhat higher than intermediate. The latter might be explained by the contribution of species-rich taxonomic groups that peaked at the maximal extent of disturbance, such as herbaceous plants. However, some disturbance-intolerant taxonomic groups, such as soil fauna and molluscs, were associated with low disturbance extent. Hence, attention should be paid not only to the peak in total diversity, but also to certain taxonomic groups and species of conservation concern.

Several recent studies have suggested that disturbance regimes of moderate extent and severity were predominant in temperate and boreal forests in the past (Stueve *et al.*, 2011; Nagel *et al.*, 2017; Čada *et al.*, 2020). Assuming that local species are adapted to historical disturbance regimes, such regimes could be expected to support the highest levels of biodiversity. However, disturbance severity can be highly variable even in the case of catastrophic disturbances, so that considerable amounts of the area can persist relatively unchanged. For example, extensive fires in Yellowstone National Park in 1988, covering thousands of square kilometres, resulted in heterogeneous burn-severity mosaics across which most of the burned area was <200 m from unburned forest edges (Turner *et al.*, 1994). Analyses of the 2019/2020 megafires in south-eastern Australia also showed that despite being greater in extent, these were not proportionally more severe than previous fires in the same area (Collins *et al.*, 2021). Nevertheless, the maximum size of high-severity patches was positively correlated with the extent of burned area. Moreover, the areas affected by megafires are becoming more homogenous, due to decreasing heterogeneity of unburnt patches as shown by recent analysis (Mackey *et al.*, 2021). Therefore, current large and extreme fires can have substantial impacts on ecosystems due to their proportionally higher burn severity and homogenising effect on forest landscapes.

Most study sites included in our synthesis were in temperate and boreal forests and investigated the effects of wildfires or windstorms; studies from other biomes and other disturbance types, matching our selection criteria, were noticeably less common. Therefore, the effects of disturbance on

biodiversity reported here mainly reflect these regions and disturbance types. The biases in the geographical distribution of disturbance types are partly natural, because certain disturbance types are more common in specific ecosystems. For example, wildfires, while being common disturbances in boreal forests, rarely occur naturally in tropical rainforests. Increasingly severe droughts induced by the El Niño Southern Oscillation (ENSO), however, have been associated with more frequent large-scale forest fire events in tropical rainforests in Asia and South America (Alencar, Solorzano & Nepstad, 2004; Gu *et al.*, 2020). Future studies investigating various types of disturbances, especially in tropical areas, would provide a valuable addition to our current knowledge.

### (3) The role of time since disturbance

In addition to revealing the effect of disturbance severity, our meta-analysis identified a temporal change in the effect of disturbance on  $\alpha$ -diversity. Consistent with previous studies investigating post-disturbance successional change, our results suggested a hump-shaped relationship between time and biodiversity (Yeboah & Chen, 2016; Hilmers *et al.*, 2018). According to our results, species richness in disturbed habitats tends to be similar to undisturbed habitat during the first years after disturbance. Thereafter, the species richness of disturbed habitats starts to increase, possibly because of colonisation through dispersal from surrounding habitats, as well as population expansion of species that survived the disturbance but remained undetected in early post-disturbance surveys because of their rarity or small size (Banks *et al.*, 2017). We found that the overall species richness in a disturbed forest reaches a plateau around 10 years after disturbance. The subsequent addition of competitive species may exclude early successional species so that species richness does not increase at the following stages while it may increase again in old-growth forests (Hilmers *et al.*, 2018).

Our results also showed that the disturbance extent corresponding to maximal diversity can increase with time since disturbance. More extensive areas of disturbed forest possibly provide suitable conditions for disturbance-dependent communities for longer periods. However, early post-disturbance conditions can be temporary and the long-term persistence of such habitats in the landscape may require further disturbances. The disturbance extent that would maintain maximal biodiversity at a greater timeframe than that covered by our study may be significantly lower than 60%. This is because the continuous presence of early successional disturbed forest covering more than half of an entire forest landscape would be the outcome of very high rates of disturbance and would probably result in drastic structural homogenisation (e.g. the loss of large old trees) and the loss of late-successional forest (Senf *et al.*, 2020). The availability of late-successional stages in a landscape is crucial for a variety of specialist taxa of late-successional forests maintaining diversity through the admixture of late-successional species (Bell, Lechowicz & Waterway, 2000; Komonen, 2001; Zhang, Kissling & He, 2013; Solar *et al.*, 2015; Watson *et al.*, 2018). Moreover, the structure and composition of

early successional forests is affected by pre-disturbance age and condition (Franklin *et al.*, 2000; Donato *et al.*, 2012; Lindenmayer *et al.*, 2019), and increasingly by anthropogenic post-disturbance management such as salvage logging (Leverkus *et al.*, 2018a). Therefore, for the persistence of high biodiversity, it is important that natural disturbances contribute to the successional heterogeneity of forest landscapes.

## V. CONCLUSIONS

- (1) We found an overall neutral effect of natural disturbance on forest  $\alpha$ -diversity, most probably due to highly variable responses of different taxonomic groups. Groups benefiting from disturbances include organisms that prefer conditions associated with open canopies. Conversely, groups associated with a ground-dwelling lifestyle and/or preference for shady conditions tended to be negatively impacted by disturbance.
- (2) The degree of disturbance severity maximising  $\alpha$ -diversity at the stand scale was similar to the degree of disturbance extent maximising the diversity in a proportional mixture of disturbed and undisturbed forests at the landscape scale, with both results suggesting that moderate disturbances support the highest levels of overall biodiversity.
- (3) Species richness reached its highest values around 10 years after the disturbance, highlighting the importance of time since disturbance in moderating the effect of disturbance on  $\alpha$ -diversity.
- (4) Natural disturbance significantly affected the species composition of six taxonomic groups (ants, birds, epigeic arachnids, epixylic lichens, and herbaceous and woody plants). By contrast, disturbance type and severity, forest biome, and time since disturbance did not modulate  $\beta$ -diversity.
- (5) Future studies on the effects of natural disturbances on biodiversity, especially focusing on rare species, less frequently studied disturbance types, and the tropical biome, could provide valuable novel information.

## VI. ACKNOWLEDGEMENTS

A. B. L. acknowledges grant LRB20/1002 from the British Ecological Society and grant B-FQM-366-UGR20 from Junta de Andalucía/FEDER. A. F. R. thanks Centro ANID Basal FB210015 (CENAMAD) and grant DIUFRO DI20-0066, Dirección de Investigación Universidad de La Frontera. C. H. G. is grateful to Stanlee Miller for his contribution to the study by Greenberg & Miller (2004). J. B. R. A. thanks the Community Ecology Lab (<https://www.fbaccaro-ecolab.com/>) and the Laboratório de Sistemática e Ecologia de Coleoptera of the National Institute for Amazonian Research (LASEC/INPA). D. M. M. and J. B. R. A. thank the Forest Management Laboratory of the National Institute for Amazonian Research (LMF/INPA) for logistic support. D. M. M. was supported by the German Federal Ministry

of Education and Research (BMBF), and the Max Planck Society (MPG). The study site of D. M. M. and J. B. R. A. is supported by the INCT Madeiras da Amazônia and the ATTO Project, which is funded by the BMBF (contracts 01LB1001A and 01LK1602A), the Brazilian Ministry of Science, Technology and Innovation (MCTI/FINEP, contract 01.11.01248.00) and the MPG. K. B. Go. and A. S. Z. thank Russian Science Foundation (project No 21-14-00227) for facilitating soil invertebrate data collection at the sites in European Russia. M. A. was supported by LTC 20058 Ministry of Education, Youth and Sports, Czech Republic and by long-term research development project of the Czech Academy of Sciences RVO67985939. M.-L. V. and S. T. were supported by the project TH 2218/5-1. S. G. P. was supported by NERC project GR3/11743. T. P. and M. L. were supported by the Greek project POL-AEGIS, Program THALES, grant MIS 376737.

We thank Jari Kouki for his comments and help with using the data published in Salo & Kouki (2018) and Salo, Domisch & Kouki (2019). We also thank an anonymous reviewer for their comments. Open Access funding enabled and organized by Projekt DEAL.

## VII. DATA AVAILABILITY STATEMENT

Original raw data will be made available by data owners on reasonable request.

## VIII. REFERENCES

- References identified with an asterisk (\*) provided data used in the meta-analysis; a dagger symbol (†) indicates additional references used only in the supporting information.
- \*ABELLA, S. R. & FORNWALT, P. J. (2015). Ten years of vegetation assembly after a North American mega fire. *Global Change Biology* **21**, 789–802.
  - \*ADÁMEK, M., HADINCOVÁ, V. & WILD, J. (2016). Long-term effect of wildfires on temperate *Pinus sylvestris* forests: vegetation dynamics and ecosystem resilience. *Forest Ecology and Management* **380**, 285–295.
  - \*ALENCAR, J. B. R., DA FONSECA, C. R. V., MARRA, D. M. & BACCARO, F. B. (2021). Windthrows promote higher diversity of saproxylic beetles (Coleoptera: Passalidae) in a Central Amazon forest. *Insect Conservation and Diversity* **15**, 1–8.
  - ALENCAR, A. A. C., SOLORIZANO, L. A. & NEPSTAD, D. C. (2004). Modeling forest understory fires in an eastern Amazonian landscape. *Ecological Applications* **14**, S139–S149.
  - ANDERSON, M. J. (2017). Permutational multivariate analysis of variance (PERMANOVA). In *Wiley StatsRef: Statistics Reference Online* (eds N. BALAKRISHNAN, T. COLTON, B. EVERITT, W. PIEGORSCH, F. RUGGERI and J. L. TEUGELS). <https://doi.org/10.1002/9781118445112.stat07841>.
  - BANKS, S. C., MCBURNEY, L., BLAIR, D., DAVIES, I. D. & LINDENMAYER, D. B. (2017). Where do animals come from during post-fire population recovery? Implications for ecological and genetic patterns in post-fire landscapes. *Ecography* **40**, 1325–1338.
  - \*BARBER, N. A. & WIDICK, W. L. (2017). Localized effects of tornado damage on ground beetle communities and vegetation in a forested preserve. *Natural Areas Journal* **37**, 489–496.
  - BATES, D., MAECHLER, M., BOLKER, B. & WALKER, S. (2015). Fitting linear mixed-effects models using lme4. *Journal of Statistical Software* **67**, 1–48.
  - BELL, G., LECHOWICZ, M. J. & WATERWAY, M. J. (2000). Environmental heterogeneity and species diversity of forest sedges. *Journal of Ecology* **88**, 67–87.
  - \*BEUDERT, B., BÄSSLER, C., THORN, S., NOSS, R., SCHRÖDER, B., DIEFFENBACH-FRIES, H., FOULLOIS, N. & MÜLLER, J. (2015). Bark beetles increase biodiversity while maintaining drinking water quality. *Conservation Letters* **8**, 272–281.

- BONGERS, F., POORTER, L., HAWTHORNE, W. D. & SHEIL, D. (2009). The intermediate disturbance hypothesis applies to tropical forests, but disturbance contributes little to tree diversity. *Ecology Letters* **12**, 798–805.
- \*BOWD, E. J., LINDENMAYER, D. B., BANKS, S. C. & BLAIR, D. P. (2018). Logging and fire regimes alter plant communities. *Ecological Applications* **28**, 826–841.
- \*BURKLE, L. A., MYERS, J. A., BELOTE, R. T. & PETERS, D. P. C. (2015). Wildfire disturbance and productivity as drivers of plant species diversity across spatial scales. *Ecosphere* **6**, 1–14.
- ČADA, V., TROTSIUK, V., JANDA, P., MIKOLÁŠ, M., BAČE, R., NAGEL, T. A., MORRISSEY, R. C., TEPLY, A. J., VOSTAREK, O., BEGOVIĆ, K., CHASKOVSKYY, O., DUŠÁTKO, M., KAMENIAR, O., KOZÁK, D., LÁBUSOVÁ, J., ET AL. (2020). Quantifying natural disturbances using a large-scale dendrochronological reconstruction to guide forest management. *Ecological Applications* **30**, 1–13.
- CARBONE, L. M., TAVELLA, J., PAUSAS, J. G. & AGUILAR, R. (2019). A global synthesis of fire effects on pollinators. *Global Ecology and Biogeography* **28**, 1487–1498.
- †CHAMBERS, J. Q., HIGUCHI, N. & SCHIMEL, J. P. (1998). Ancient trees in Amazonia. *Nature* **391**, 135–136.
- CHAO, A., COLWELL, R. K., GOTELLI, N. J. & THORN, S. (2019). Proportional mixture of two rarefaction/extrapolation curves to forecast biodiversity changes under landscape transformation. *Ecology Letters* **22**, 1913–1922.
- CHAO, A., GOTELLI, N. J., HSIEH, T. C., SANDER, E. L., MA, K. H., COLWELL, R. K. & ELLISON, A. M. (2014). Rarefaction and extrapolation with Hill numbers: a framework for sampling and estimation in species diversity studies. *Ecological Monographs* **84**, 45–67.
- CHAO, A., KUBOTA, Y., ZELENY, D., CHIU, C.-H., LI, C.-F., KUSUMOTO, B., YASUHARA, M., THORN, S., WEI, C.-L., COSTELLO, M. J. & COLWELL, R. K. (2020). Quantifying sample completeness and comparing diversities among assemblages. *Ecological Research* **35**, 292–314.
- CHASE, J. M. & MYERS, J. A. (2011). Disentangling the importance of ecological niches from stochastic processes across scales. *Philosophical Transactions of the Royal Society B* **366**, 2351–2363.
- \*CHERGUL, B., FAHD, S. & SANTOS, X. (2019). Are reptile responses to fire shaped by forest type and vegetation structure? Insights from the Mediterranean basin. *Forest Ecology and Management* **437**, 340–347.
- CHESSON, P. & HUNTLY, N. (1997). The roles of harsh and fluctuating conditions in the dynamics of ecological communities. *The American Naturalist* **150**, 519–553.
- \*CHOI, C. Y., LEE, E. J., NAM, H. Y., LEE, W. S. & LIM, J. H. (2014). Temporal changes in the breeding bird community caused by post-fire treatments after the Samcheok forest fire in Korea. *Landscape and Ecological Engineering* **10**, 203–214.
- \*CLEARY, D. F. R. & MOOERS, A. (2006). Burning and logging differentially affect endemic vs. widely distributed butterfly species in Borneo. *Diversity and Distributions* **12**, 409–416.
- CLEMENTS, F. E. (1916). *Plant Succession: An Analysis of the Development of Vegetation*. Carnegie Institute of Washington, Washington, DC.
- COLLINS, L., BRADSTOCK, R. A., CLARKE, H., CLARKE, M. F., NOLAN, R. H. & PENMAN, T. D. (2021). The 2019/2020 mega-fires exposed Australian ecosystems to an unprecedented extent of high-severity fire. *Environmental Research Letters* **16**, 1–14.
- CONNELL, J. H. (1978). Diversity in tropical rain forests and coral reefs. *Science* **199**, 1302–1310.
- \*ČUCHTA, P., MIKLISOVÁ, D. & KOVÁČ, Ľ. (2019). The succession of soil Collembola communities in spruce forests of the High Tatras Mountains five years after a windthrow and clear-cut logging. *Forest Ecology and Management* **433**, 504–513.
- CUMMING, S. G. (2005). Effective fire suppression in boreal forests. *Canadian Journal of Forest Research* **35**, 772–786.
- \*DAVIS, T. S., RHOADES, P. R., MANN, A. J. & GRISWOLD, T. (2020). Bark beetle outbreak enhances biodiversity and foraging habitat of native bees in alpine landscapes of the southern Rocky Mountains. *Scientific Reports* **10**, 1–14.
- \*DECHNIK-VÁZQUEZ, Y. A., MEAVE, J. A., PÉREZ-GARCÍA, E. A., GALLARDO-CRUZ, J. A. & ROMERO-ROMERO, M. A. (2016). The effect of treefall gaps on the understorey structure and composition of the tropical dry forest of Nizanda, Oaxaca, Mexico: implications for forest regeneration. *Journal of Tropical Ecology* **32**, 89–106.
- \*DODDS, K. J., DIGIROLOMO, M. F. & FRAVER, S. (2019). Response of bark beetles and woodborers to tornado damage and subsequent salvage logging in northern coniferous forests of Maine, USA. *Forest Ecology and Management* **450**, 1–14.
- DONATO, D. C., CAMPBELL, J. L. & FRANKLIN, J. F. (2012). Multiple successional pathways and precocity in forest development: can some forests be born complex? *Journal of Vegetation Science* **23**, 576–584.
- \*DOWNING, W. M., KRAWCHUK, M. A., COOP, J. D., MEIGS, G. W., HAIRE, S. L., WALKER, R. B., WHITMAN, E., CHONG, G., MILLER, C. & TORTORELLI, C. (2020). How do plant communities differ between fire refugia and fire-generated early-seral vegetation? *Journal of Vegetation Science* **31**, 26–39.
- \*DUELLI, P., OBRIST, M. K. & WERMELINGER, B. (2002). Windthrow-induced changes in faunistic biodiversity in alpine spruce forests. *Forest Snow and Landscape Research* **77**, 117–131.
- FAHRIG, L. (2020). Why do several small patches hold more species than few large patches? *Global Ecology and Biogeography* **29**, 615–628.
- FOX, J. W. (2013). The intermediate disturbance hypothesis should be abandoned. *Trends in Ecology and Evolution* **28**, 86–92.
- FRANCO-MANCHÓN, I., SALO, K., ORIA-DE-RUEDA, J. A., BONET, J. A. & MARTÍN-PINTO, P. (2019). Are wildfires a threat to fungi in European *Pinus* forests? A case study of boreal and Mediterranean forests. *Forests* **10**, 1–12.
- FRANKLIN, J. F., LINDENMAYER, D. B., MACMAHON, J. A., MCKEE, A., MAGNUSON, J., PERRY, D. A., WAIDE, R. & FOSTER, D. (2000). Threads of continuity. *Conservation Biology* **1**, 8–17.
- \*FUENTES-RAMÍREZ, A., SALAS-ELJATIB, C., GONZÁLES, M. E., URRUTIA-ESTRADA, J., ARROYO-VARGAS, P. & SANTIBÁNEZ, P. (2020). Initial response of understorey vegetation and tree regeneration to a mixed-severity fire in old-growth *Araucaria-Nothofagus* forests. *Applied Vegetation Science* **23**, 210–222.
- GAMFELDT, L., SNÄLL, T., BAGCHI, R., JONSSON, M., GUSTAFSSON, L., KJELLANDER, P., RUIZ-JAEN, M. C., FRÖBERG, M., STENDAHL, J., PHILIPSON, C. D., MIKUSIŃSKI, G., ANDERSSON, E., WESTERLUND, B., ANDRÉN, H., MOBERG, F., ET AL. (2013). Higher levels of multiple ecosystem services are found in forests with more tree species. *Nature Communications* **4**, 1–8.
- \*GANDHI, K. J. K., GILMORE, D. W., KATOVICH, S. A., MATTSO, W. J., ZASADA, J. C. & SEYBOLD, S. J. (2008). Catastrophic windstorm and fuel-reduction treatments alter ground beetle (Coleoptera: Carabidae) assemblages in a North American sub-boreal forest. *Forest Ecology and Management* **256**, 1104–1123.
- \*GEORGIEV, K., BEUDERT, B., BÄSSLER, C., FELDHAAR, H., HEIBL, C., KARASCH, P., MÜLLER, J., PERLÍK, M., WEISS, I. & THORN, S. (2021). Forest disturbance and salvage logging have neutral long-term effects on drinking water quality but alter biodiversity. *Forest Ecology and Management* **495**, 119354.
- GEORGIEV, K. B., CHAO, A., CASTRO, J., CHEN, Y.-H., CHOI, C.-Y., FONTAINE, J. B., HUTTO, R. L., LEE, E.-J., MÜLLER, J., ROST, J., ŽMIHORSKI, M. & THORN, S. (2020). Salvage logging changes the taxonomic, phylogenetic and functional trajectories of forest bird communities. *Journal of Applied Ecology* **57**, 1103–1112.
- \*GONGALSKY, K. B., ZAITSEV, A. S., KOROBUSHKIN, D. I., SAIFUTDINOV, R. A., BUTENKO, K. O., DE VRIES, F. T., EKSCHMITT, K., DEGTYAREV, M. I., GORBUNOVA, A. Y., KOSTINA, N. V., RAKHLEVA, A. A., SHAKHAB, S. V., YAZRIKOVA, T. E., WOLTERS, V. & BARDGETT, R. D. (2021). Forest fire induces short-term shifts in soil food webs with consequences for carbon cycling. *Ecology Letters* **24**, 438–450.
- \*GREENBERG, C. H. (2001). Response of reptile and amphibian communities to canopy gaps created by wind disturbance in the southern Appalachians. *Forest Ecology and Management* **148**, 135–144.
- \*GREENBERG, C. H. & LANHAM, J. D. (2001). Breeding bird assemblages of hurricane-created gaps and adjacent closed canopy forest in the southern Appalachians. *Forest Ecology and Management* **154**, 251–260.
- \*GREENBERG, C. H. & MILLER, S. (2004). Soricid response to canopy gaps created by wind disturbance in the southern Appalachians. *Southeastern Naturalist* **3**, 715–732.
- GRIME, J. P. (1973). Competitive exclusion in herbaceous vegetation. *Nature* **242**, 344–347.
- GU, Y., LIU, H., TRAORÉ, D. D. & HUANG, C. (2020). ENSO-related droughts and ISM variations during the last millennium in tropical Southwest China. *Climate Dynamics* **54**, 649–659.
- HALME, P., TOIVANEN, T., HONKANEN, M., KOTIAHO, J. S., MÖNKKÖNEN, M. & TIMONEN, J. (2010). Flawed meta-analysis of biodiversity effects of forest management. *Conservation Biology* **24**, 1154–1156.
- HASTINGS, A. (1989). Disturbance, coexistence, history, and competition for space. *Theoretical Population Biology* **18**, 363–373.
- HEDGES, L. V., GUREVITCH, J. & CURTIS, P. S. (1999). The meta-analysis of response ratios in experimental ecology. *Ecology* **80**, 1150–1156.
- HEDWALL, P.-O. & MIKUSIŃSKI, G. (2016). Sprucification in protected forests: myth or veracity? Clues from 60 yrs survey data. *Applied Vegetation Science* **19**, 317–380.
- HILL, M. O. (1973). Diversity and evenness: a unifying notation and its consequences. *Ecology* **54**, 427–432.
- HILL, S. L. L., ARNELL, A., MANEY, C., BUTCHART, S. H. M., HILTON-TAYLOR, C., CICIARELLI, C., DAVIS, C., DINERSTEIN, E., PURVIS, A. & BURGESS, N. D. (2019). Measuring forest biodiversity status and changes globally. *Frontiers in Forests and Global Change* **2**, 1–11.
- HILMERS, T., FRIESS, N., BÄSSLER, C., HEURICH, M., BRANDL, R., PRETZSCH, H., SEIDL, R. & MÜLLER, J. (2018). Biodiversity along temperate forest succession. *Journal of Applied Ecology* **55**, 2756–2766.
- HSIEH, T. C., MA, K. H. & CHAO, A. (2016). iNEXT: an R package for rarefaction and extrapolation of species diversity (Hill numbers). *Methods in Ecology and Evolution* **7**, 1451–1456.
- HURLBERT, S. H. (1984). Pseudoreplication and the design of ecological field experiments. *Ecological Monographs* **54**, 187–211.
- \*HYLANDER, K. (2011). The response of land snail assemblages below aspens to forest fire and clear-cutting in Fennoscandian boreal forests. *Forest Ecology and Management* **261**, 1811–1819.

- JONES, G. M. & TINGLEY, M. W. (2021). Pyrodiversity and biodiversity: a history, synthesis and outlook. *Diversity and Distributions* **28**, 386–403. <https://doi.org/10.1111/ddi.13280>.
- KEELEY, J. E. (2009). Fire intensity, fire severity and burn severity: a brief review and suggested usage. *International Journal of Wildland Fire* **18**, 116–126.
- KING, W. (1685). On the bogs and loughs of Ireland. *Philosophical Transactions of the Royal Society of London* **15**, 948–960.
- KOMONEN, A. (2001). Structure of insect communities inhabiting old-growth forest specialist bracket fungi. *Ecological Entomology* **26**, 63–75.
- KORTMANN, M., MÜLLER, J. C., BAIER, R., BÄSSLER, C., BUSE, J., CHOLEWIŃSKA, O., FÖRSCHLER, M. I., GEORGIEV, K. B., HILSZCZAŃSKI, J., JAROSZEWICZ, B., JAWORSKI, T., KAUFMANN, S., KUJPER, D., LORZ, J., LOTZ, A., *ET AL.* (2021). Ecology versus society: impacts of bark beetle infestations on biodiversity and restorativeness in protected areas of Central Europe. *Biological Conservation* **254**, 2–9.
- \*KRNO, I., LÁNCZOS, T. & ŠPORKA, F. (2015). Windstorm disturbance effects on mountain stream ecosystems and the Plecoptera assemblages. *Biologia (Poland)* **70**, 1215–1227.
- \*KURTH, V. J., FRANSIOLI, N., FULÉ, P. Z., HART, S. C. & GEHRING, C. A. (2013). Stand-replacing wildfires alter the community structure of wood-inhabiting fungi in southwestern ponderosa pine forests of the USA. *Fungal Ecology* **6**, 192–204.
- LAJEUNESSE, M. J. (2011). On the meta-analysis of response ratios for studies with correlated and multi-group designs. *Ecology* **92**, 2049–2055.
- LAZARINA, M., DEVALEZ, J., NEOKOSMIDIS, L., SGARDELIS, S. P., KALLIMANIS, A. S., TSCHULIN, T., TSALKATIS, P., KOURTIDOU, M., MIZERAKIS, V., NAKAS, G., PALAIOLOGOU, P., KALABOKIDIS, K., VUJIC, A. & PETANIDOU, T. (2019). Moderate fire severity is best for the diversity of most of the pollinator guilds in Mediterranean pine forest. *Ecology* **100**, e02615.
- \*LAZARINA, M., SGARDELIS, S. P., TSCHULIN, T., DEVALEZ, J., MIZERAKIS, V., KALLIMANIS, A. S., PAPAKONSTANTINO, S., KYRIAZIS, T. & PETANIDOU, T. (2017). The effect of fire history in shaping diversity patterns of flower-visiting insects in post-fire Mediterranean pine forests. *Biodiversity and Conservation* **26**, 115–131.
- \*LAZARINA, M., SGARDELIS, S. P., TSCHULIN, T., KALLIMANIS, A. S., DEVALEZ, J. & PETANIDOU, T. (2016). Bee response to fire regimes in Mediterranean pine forests: the role of nesting preference, trophic specialization, and body size. *Basic and Applied Ecology* **17**, 308–320.
- LEVERKUS, A. B., LINDENMAYER, D. B., THORN, S. & GUSTAFSSON, L. (2018a). Salvage logging in the world's forests: interactions between natural disturbance and logging need recognition. *Global Ecology and Biogeography* **27**, 1140–1154.
- LEVERKUS, A. B., REY BENAYAS, J. M., CASTRO, J., BOUCHER, D., BREWER, S., COLLINS, B. M., DONATO, D., FRAVER, S., KISHCHUK, B. E., LEE, E. J., LINDENMAYER, D. B., LINGUA, E., MACDONALD, E., MARZANO, R., RHOADES, C. C., *ET AL.* (2018b). Salvage logging effects on regulating and supporting ecosystem services—a systematic map. *Canadian Journal of Forest Research* **48**, 983–1000.
- \*LINDENMAYER, D. B., MCBURNEY, L., BLAIR, D., WOOD, J. & BANKS, S. C. (2018). From unburnt to salvage logged: quantifying bird responses to different levels of disturbance severity. *Journal of Applied Ecology* **55**, 1626–1636.
- LINDENMAYER, D. B. & TAYLOR, C. (2020). New spatial analyses of Australian wildfires highlight the need for new fire, resource, and conservation policies. *Proceedings of the National Academy of Sciences of the United States of America* **117**, 12481–12485.
- LINDENMAYER, D. B., WESTGATE, M. J., SCHEELE, B. C., FOSTER, C. N. & BLAIR, D. P. (2019). Key perspectives on early successional forests subject to stand-replacing disturbances. *Forest Ecology and Management* **454**, 1–10.
- MACKEY, B., LINDENMAYER, D., NORMAN, P., TAYLOR, C. & GOULD, S. (2021). Are fire refugia less predictable due to climate change? *Environmental Research Letters* **16**, 114028.
- MACKEY, R. L. & CURRIE, D. J. (2001). The diversity-disturbance relationship: is it generally strong and peaked? *Ecology* **82**, 3479–3492.
- MARTIN, R. E. & SAPSIS, D. B. (1992). Fires as agents of biodiversity: pyrodiversity promotes biodiversity. In *Proceedings of the Symposium on Biodiversity of Northwestern California* (eds R. C. HARRIS, D. C. ERMAN and H. M. KERNER), pp. 150–157. Wildland Resources Center, University of California, Berkeley.
- \*MARTÍN-PINTO, P., VAQUERIZO, H., PEÑALVER, F., OLAIZOLA, J. & ORIA-DE-RUEDA, J. A. (2006). Early effects of a wildfire on the diversity and production of fungal communities in Mediterranean vegetation types dominated by *Cistus ladanifer* and *Pinus pinaster* in Spain. *Forest Ecology and Management* **225**, 296–305.
- MCDARLE, B. H. & ANDERSON, M. J. (2001). Fitting multivariate models to community data: a comment on distance-based redundancy analysis. *Ecology* **82**, 290–297.
- \*MORETTI, M. & BARBALAT, S. (2004). The effects of wildfires on wood-eating beetles in deciduous forests on the southern slope of the Swiss Alps. *Forest Ecology and Management* **187**, 85–103.
- \*MORETTI, M., CONEDERA, M., DUELLI, P. & EDWARDS, P. J. (2002). The effects of wildfire on ground-active spiders in deciduous forests on the Swiss southern slope of the Alps. *Journal of Applied Ecology* **39**, 321–336.
- \*MORETTI, M., DE BELLO, F., ROBERTS, S. P. M. & POTTS, S. G. (2009). Taxonomical vs. functional responses of bee communities to fire in two contrasting climatic regions. *Journal of Animal Ecology* **78**, 98–108.
- \*MORETTI, M., DE CÀCERES, M., PRADILLA, C., OBRIST, M. K., LEGENDRE, P. & DUELLI, P. (2010). Fire-induced taxonomic and functional changes in saproxylic beetle communities in fire sensitive regions. *Ecography* **33**, 760–771.
- MORETTI, M. & LEGG, C. (2009). Combining plant and animal traits to assess community functional responses to disturbance. *Ecography* **32**(2), 299–309.
- MORETTI, M., OBRIST, M. K. & DUELLI, P. (2004). Arthropod biodiversity after forest fires: winners and losers in the winter fire regime of the southern Alps. *Ecography* **27**(2), 173–186.
- NAGEL, T. A., MIKAC, S., DOLINAR, M., KLOPCIC, M., KEREN, S., SVOBODA, M., DIACI, J., BONCINA, A. & PAULIC, V. (2017). The natural disturbance regime in forests of the Dinaric Mountains: a synthesis of evidence. *Forest Ecology and Management* **388**, 29–42.
- NELSON, C. R., HALPERN, C. B. & AGEE, J. K. (2008). Thinning and burning result in low-level invasion by nonnative plants but neutral effects on natives. *Ecological Applications* **18**, 762–770.
- OKSANEN, J., BLANCHET, F. G., FRIENDLY, M., KINDT, R., LEGENDRE, P., MCGILN, D., MINCHIN, P. R., O'HARA, R. B., SIMPSON, G. L., SOLYMS, P., STEVENS, M. H. H., SZOEC, E. & WAGNER, H. (2019). Vegan: community ecology package. R package version 2.5-6. Electronic file available at <https://CRAN.R-project.org/package=vegan>. Accessed August 01, 2020.
- OLSON, D. M., DINERSTEIN, E., WIKRAMANAYAKE, E. D., BURGESS, N. D., POWELL, G. V. N., UNDERWOOD, E. C., D'AMICO, J. A., ITOUA, I., STRAND, H. E., MORRISON, J. C., LOUCKS, C. J., ALLNUTT, T. F., RICKETTS, T. H., KURA, Y., LAMOREUX, J. F., *ET AL.* (2001). Terrestrial ecoregions of the world: a new map of life on Earth. *Bioscience* **51**, 933–938.
- †PAGE, M. J., MCKENZIE, J. E., BOSSUYT, P. M., BOUTRON, I., HOFFMANN, T. C., MULROW, C. D., SHAMSEER, L., TETZLAFF, J. M., AKI, E. A., BRENNAN, S. E., CHOU, R., GLANVILLE, J., GRIMSHAW, J. M., HRÓBJARTSSON, A., LALU, M. M., *ET AL.* (2021). The PRISMA 2020 statement: an updated guideline for reporting systematic reviews. *Research Methods & Reporting* **372**, n71.
- PAQUETTE, A. & MESSIER, C. (2011). The effect of biodiversity on tree productivity: from temperate to boreal forests. *Global Ecology and Biogeography* **20**, 170–180.
- \*PARRO, K., KÖSTER, K., JÖGISTE, K. & VODDE, F. (2009). Vegetation dynamics in a fire damaged forest area: the response of major ground vegetation species. *Baltic Forestry* **15**, 206–215.
- PAUSAS, J. G. (2019). Generalized fire response strategies in plants and animals. *Oikos* **128**, 147–153.
- PAUSAS, J. G., BRADSTOCK, R. A., KEITH, D. A., KEELEY, J. E. & THE GCTE (GLOBAL CHANGE OF TERRESTRIAL ECOSYSTEMS) FIRE NETWORK (2004). Plant functional traits in relation to fire in crown-fire ecosystems. *Ecology* **84**, 1085–1100.
- PICKETT, S. T. A. & WHITE, P. S. (1985). *The Ecology of Natural Disturbance and Patch Dynamics*. Academic Press, New York.
- PONISIO, L. C., WILKIN, K., M'GONIGLE, L. K., KULHANKE, K., COOK, L., THORP, R., GRISWOLD, T. & KREMEN, C. (2016). Pyrodiversity begets plant-pollinator community diversity. *Global Change Biology* **22**, 1794–1808.
- \*PONS, P. & WENDENBURG, C. (2005). The impact of fire and forest conversion into savanna on the bird communities of West Madagascan dry forests. *Animal Conservation* **8**, 183–193.
- \*POTTS, S. G., VULLIAMY, B., DAFNI, A., NE, G., TOOLE, C. O., ROBERTS, S. & WILLMER, P. (2003). Response of plant-pollinator communities to fire: changes in diversity, abundance and floral reward. *Oikos* **101**, 103–112.
- PROCTOR, E., NOL, E., BURKE, D. & CRINS, W. J. (2012). Responses of insect pollinators and understory plants to silviculture in northern hardwood forests. *Biodiversity and Conservation* **21**, 1703–1740.
- PULLIN, A. S. & STEWART, G. B. (2006). Guidelines for systematic review in conservation and environmental management. *Conservation Biology* **20**, 1647–1656.
- PULSFORD, S. A., LINDENMAYER, D. B. & DRISCOLL, D. A. (2016). A succession of theories: purging redundancy from disturbance theory. *Biological Reviews* **91**, 148–167.
- R CORE TEAM (2021). *R: A Language and Environment for Statistical Computing*. R Foundation for Statistical Computing, Vienna. Electronic file available at <https://www.R-project.org/>. Accessed August 01, 2020.
- RAMOS-FABIEL, M. A., PÉREZ-GARCÍA, E. A., GONZÁLEZ, E. J., YÁÑEZ-ORDÓÑEZ, O. & MEAVE, J. A. (2018). Successional dynamics of the bee community in a tropical dry forest: insights from taxonomy and functional ecology. *Biotropica* **51**, 62–74.
- \*RATCHFORD, J. S., WITTMAN, S. E., JULES, E. S., ELLISON, A. M., GOTELLI, N. J. & SANDERS, N. J. (2005). The effects of fire, local environment and time on ant assemblages in fens and forests. *Diversity and Distributions* **11**, 487–497.
- REICE, S. R., WISSMAR, R. C. & NAIMAN, R. J. (1990). Disturbance regimes, resilience, and recovery of animal communities and habitats in lotic ecosystems. *Environmental Management* **14**, 647–659.



- RHOADES, P. R., DAVIS, T. S., TINKHAM, W. T. & HOFFMANN, C. M. (2018). Effects of seasonality, forest structure, and understory plant richness on bee community assemblage in a Southern Rocky Mountain mixed conifer forest. *Annals of the Entomological Society of America* **111**, 278–284.
- \*RICHTER, C., REJMÁNEK, M., MILLER, J. E. D., WELCH, K. R., WEEKS, J. M. & SAFFORD, H. (2019). The species diversity × fire severity relationship is hump-shaped in semiarid yellow pine and mixed conifer forests. *Ecosphere* **10**, 1–16.
- RODRÍGUEZ, A. & KOUKI, J. (2017). Disturbance-mediated heterogeneity drives pollinator diversity in boreal managed forest ecosystems. *Ecological Applications* **27**, 589–602.
- RUBENE, D., SCHROEDER, M. & RANIUS, T. (2015). Diversity patterns of wild bees and wasps in managed boreal forests: effects of spatial structure, local habitat and surrounding landscape. *Biological Conservation* **184**, 201–208.
- \*SAIFUTDINOV, R. A., GONGALSKY, K. B. & ZAITSEV, A. S. (2018). Evidence of a trait-specific response to burning in springtails (Hexapoda: Collembola) in the boreal forests of European Russia. *Goderma* **332**, 173–179.
- \*SALO, K., DOMISCH, T. & KOUKI, J. (2019). Forest wildfire and 12 years of post-disturbance succession of saprotrophic macrofungi (Basidiomycota, Ascomycota). *Forest Ecology and Management* **451**, 117454.
- \*SALO, K. & KOUKI, J. (2018). Severity of forest wildfire had a major influence on early successional ectomycorrhizal macrofungi assemblages, including edible mushrooms. *Forest Ecology and Management* **415–416**, 70–84.
- \*SCHMALHOLZ, M., HYLANDER, K. & FREGO, K. (2011). Bryophyte species richness and composition in young forests regenerated after clear-cut logging versus after wildfire and spruce budworm outbreak. *Biodiversity and Conservation* **20**, 2575–2596.
- SCHULZE, K., MALEK, Z. & VERBURG, P. H. (2019). Towards better mapping of forest management patterns: a global allocation approach. *Forest Ecology and Management* **432**, 776–785.
- SEIBOLD, S., BÄSSLER, C., BRANDL, R., BÜCHE, B., SZALLIES, A., THORN, S., ULYSHEN, M. D. & MÜLLER, J. (2016). Microclimate and habitat heterogeneity as the major drivers of beetle diversity in dead wood. *Journal of Applied Ecology* **53**, 934–943.
- SEIBOLD, S., RAMMER, W., HOTHORN, T., SEIDL, R., ULYSHEN, M. D., LÖRZ, J., CADOTTE, M. W., LINDENMAYER, D. B., ADHIKARI, Y. P., ARAGÓN, R., BAE, S., BALDRAN, P., VARANDI, H. B. M., BARLOW, J., BÄSSLER, C., ET AL. (2021). The contribution of insects to global forest deadwood decomposition. *Nature* **597**, 77–84.
- SEIDL, R., THOM, D., KAUTZ, M., MARTIN-BENITO, D., PELTONIEMI, M., VACCHIANO, G., WILD, J., ASCOLI, D., PETR, M., HONKANENIEMI, J., LEXER, M. J., TROTSIUK, V., MAIROTA, P., SVOBODA, M., FABRIKA, M., ET AL. (2017). Forest disturbances under climate change. *Nature Climate Change* **7**, 395–402.
- SENF, C., MORI, A. S., MÜLLER, J. & SEIDL, R. (2020). The response of canopy height diversity to natural disturbances in two temperate forest landscapes. *Landscape Ecology* **35**, 2101–2112.
- SOLAR, R. R., BARLOW, J., FERREIRA, J., BERENGUER, E., LEES, A. C., THOMSON, J. R., LOUZADA, J., MAUÉS, M., MOURA, N. G., OLIVEIRA, V. H. F., CHAUL, J. C. M., SCHROEDER, J. H., VIEIRA, I. C. G., MAC NALLY, R. & GARDNER, T. A. (2015). How pervasive is biotic homogenization in human-modified tropical forest landscapes? *Ecology Letters* **18**, 1108–1118.
- SOMMERFELD, A., SENF, C., BUMA, B., D'AMATO, A. W., DESPRÉS, T., DÍAZ-HORMAZÁBAL, I., FRAVER, S., FRELICH, L. E., GUTIÉRREZ, Á. G., HART, S. J., HARVEY, B. J., HE, H. S., HLÁSNY, T., HOLZ, A., KITZBERGER, T., ET AL. (2018). Patterns and drivers of recent disturbances across the temperate forest biome. *Nature Communications* **9**, 1–9.
- \*STEEL, Z. L., CAMPOS, B., FRICK, W. F., BURNETT, R. & SAFFORD, H. D. (2019). The effects of wildfire severity and pyrodiversity on bat occupancy and diversity in fire-suppressed forests. *Scientific Reports* **9**, 1–11.
- STEEL, Z. L., COLLINS, B. M., SAPSIS, D. B. & STEPHENS, S. L. (2021). Quantifying pyrodiversity and its drivers. *Proceedings of the Royal Society B* **288**, 20203202.
- STEIN, A., GERSTNER, K. & KREFT, H. (2014). Environmental heterogeneity as a universal driver of species richness across taxa, biomes and spatial scales. *Ecology Letters* **17**, 866–880.
- \*STRAND, E. K., SATTERBERG, K. L., HUDAK, A. T., BYRNE, J., KHALYANI, A. H. & SMITH, A. M. S. (2019). Does burn severity affect plant community diversity and composition in mixed conifer forests of the United States Intermountain West one decade post fire? *Fire Ecology* **15**, 25.
- STUEVE, K. M., HOBIE PERRY, C. H., NELSON, M. D., HEALEY, S. P., HILL, A. D., MOISEN, G. G., COHEN, W. B., GORMANSON, D. D. & HUANG, C. (2011). Ecological importance of intermediate windstorms rivals large, infrequent disturbances in the northern Great Lakes. *Ecosphere* **2**, 1–21.
- SWANSON, M. E., FRANKLIN, J. F., BESCHTA, R. L., CRISAFULLI, C. M., DELLASALA, D. A., HÜTTO, R. L., LINDENMAYER, D. B. & SWANSON, F. J. (2011). The forgotten stage of forest succession: early-successional ecosystems on forest sites. *Frontiers in Ecology and the Environment* **9**, 117–125.
- \*TAVELLA, J. & CAGNOLO, L. (2019). Does fire disturbance affect ant community structure? Insights from spatial co-occurrence networks. *Oecologia* **189**, 475–486.
- THOM, D. & SEIDL, R. (2015). Natural disturbance impacts on ecosystem services and biodiversity in temperate and boreal forests. *Biological Reviews* **91**, 760–781.
- THORN, S., CHAO, A., BERNHARDT-RÖRMANN, M., CHEN, Y.-H., GEORGIEV, K. B., HEIBL, C., MÜLLER, J., SCHÄFER, H. & BÄSSLER, C. (2020a). Rare species, functional groups, and evolutionary lineages drive successional trajectories in disturbed forests. *Ecology* **101**, e02949.
- THORN, S., CHAO, A., GEORGIEV, K. B., MÜLLER, J., BÄSSLER, C., CAMPBELL, J. L., CASTRO, J., CHEN, Y. H., CHOI, C. Y., COBB, T. P., DONATO, D. C., DURSKA, E., MACDONALD, E., FELDHAAR, H., FONTAINE, J. B., ET AL. (2020b). Estimating retention benchmarks for salvage logging to protect biodiversity. *Nature Communications* **11**, 1–8.
- \*THORN, S., HACKER, H. H., SEIBOLD, S., JEHL, H., BÄSSLER, C. & MÜLLER, J. (2015). Guild-specific responses of forest Lepidoptera highlight conservation-oriented forest management-implications from conifer-dominated forests. *Forest Ecology and Management* **337**, 41–47.
- \*THORN, S., WERNER, S. A. B., WOHLFAHRT, J., BÄSSLER, C., SEIBOLD, S., QUILLFELDT, P. & MÜLLER, J. (2016). Response of bird assemblages to windstorm and salvage logging—insights from analyses of functional guild and indicator species. *Ecological Indicators* **65**, 142–148.
- \*TRUCCHI, E., PITZALIS, M., ZAPPAROLI, M. & BOLOGNA, A. M. (2009). Short-term effects of canopy and surface fire on centipede (Chilopoda) communities in a semi natural Mediterranean forest. *Entomologica Fennica* **20**, 129–138.
- TURNER, M. G., HARGROVE, W. W., GARDNER, R. H. & ROMME, W. H. (1994). Effects of fire on landscape heterogeneity in Yellowstone National Park, Wyoming. *Journal of Vegetation Science* **5**, 731–742.
- VIECHTBAUER, W. (2010). Conducting meta-analyses in R with the metafor package. *Journal of Statistical Software* **36**, 1–48.
- †VIEIRA, S., TRUMBORE, S., CAMARGO, P. B., SELHORST, D., CHAMBERS, J. Q., HIGUCHI, N. & MARTINELLI, L. A. (2005). Slow growth rates of Amazonian trees: consequences for carbon cycling. *Proceedings of the National Academy of Sciences of the United States of America* **102**, 18502–18507.
- WATSON, J. E. M., EVANS, T., VENTER, O., WILLIAMS, B., TULLOCH, A., STEWART, C., THOMPSON, I., RAY, J. C., MURRAY, K., SALAZAR, A., MCALPINE, C., POTAPOV, P., WALSTON, J., ROBINSON, J. G., PAINTER, M., ET AL. (2018). The exceptional value of intact forest ecosystems. *Nature Ecology and Evolution* **2**, 599–610.
- \*WERMELINGER, B., MORETTI, M., DUELLI, P., LACHAT, T. G. B. & OBRIST, M. K. (2017). Impact of windthrow and salvage-logging on taxonomic and functional diversity of forest arthropods. *Forest Ecology and Management* **391**, 9–18.
- WERNBERG, T., COURAUDON-RÉALE, M., TUYA, F. & THOMSEN, M. (2020). Disturbance intensity, disturbance extent and ocean climate modulate kelp forest understory communities. *Marine Ecology Progress Series* **651**, 57–69.
- WHITTAKER, R. H. (1960). Vegetation of the Siskiyou mountains, Oregon and California. *Ecological Monographs* **30**, 279–338.
- WHITTAKER, R. H. (1972). Evolution and measurement of species diversity. *Taxon* **21**, 213–251.
- WOOD, S. N., PYA, N. & SAEFKEN, B. (2016). Smoothing parameter and model selection for general smooth models (with discussion). *Journal of the American Statistical Association* **111**, 1548–1575.
- WOOTTON, J. T. (1998). Effects of disturbance on species diversity: a multitrophic perspective. *The American Naturalist* **152**, 803–825.
- YEOBAH, D. & CHEN, H. Y. H. (2016). Diversity–disturbance relationship in forest landscapes. *Landscape Ecology* **31**, 981–987.
- \*ZAITSEV, A. S., GONGALSKY, K. B., PERSSON, T. & BENGTSSON, J. (2014). Connectivity of litter islands remaining after a fire and unburnt forest determines the recovery of soil fauna. *Applied Soil Ecology* **83**, 101–108.
- ZHANG, J., KISSLING, W. D. & HE, F. (2013). Local forest structure, climate and human disturbance determine regional distribution of boreal bird species richness in Alberta, Canada. *Journal of Biogeography* **40**, 1131–1142.

## IX. SUPPORTING INFORMATION

Additional supporting information may be found online in the Supporting Information section at the end of the article.

**Fig. S1.** PRISMA (Preferred Reporting Items for Systematic reviews and Meta-Analyses) diagram representing the flow of information through the decision process (i.e. the number of studies identified, rejected, and accepted).

**Table S1.** The levels of disturbance severity used in the meta-analysis for quantifying disturbance severity at each study site.

**Table S2.** Characteristics of undisturbed control forests in all study sites included in the meta-analysis.

**Fig. S2.** Differences in  $\alpha$ -diversity between naturally disturbed and undisturbed forests categorised by taxonomic group.

**Table S3.** Results of the mixed-effects meta-analysis.

**Table S4.** Results of the generalised linear mixed models analysing the natural disturbance effects on species diversity (Hill number,  $q = 0$ ) with 95% sampling coverage.

**Table S5.** Results of the generalised linear mixed models analysing the natural disturbance effects on species diversity (Hill number,  $q = 0$ ) with 95% sampling coverage (same model as in Table S4).

**Fig. S3.** Differences in diversity of all species ( $q = 0$ ) between naturally disturbed and undisturbed forests categorised by taxonomic group.

**Table S6.** Results of the generalised linear mixed models analysing the natural disturbance effects on species diversity (Hill number,  $q = 1$ ) with 95% sampling coverage.

**Table S7.** Results of the generalised linear mixed models analysing the natural disturbance effects on species diversity

(Hill number,  $q = 1$ ) with 95% sampling coverage (same model as in Table S6).

**Fig. S4.** Differences in diversity of common species ( $q = 1$ ) between naturally disturbed and undisturbed forests categorised by taxonomic group.

**Table S8.** Results of the generalised linear mixed models analysing the natural disturbance effects on species diversity (Hill number,  $q = 2$ ) with 95% sampling coverage.

**Table S9.** Results of the generalised linear mixed models analysing the natural disturbance effects on species diversity (Hill number,  $q = 2$ ) with 95% sampling coverage (same model as in Table S8).

**Fig. S5.** Differences in diversity of dominant species ( $q = 2$ ) between naturally disturbed and undisturbed forests categorised by taxonomic group.

**Table S10.** Results of linear mixed-effects model with pseudo  $F$ -value as the response variable.

**Table S11.** Results of the beta-regression with the proportion of disturbed forest corresponding to highest mixture diversity as a response variable.

(Received 15 September 2021; revised 12 May 2022; accepted 17 May 2022; published online 8 July 2022)